

## AN EXTENSION OF KESTEN'S GENERALISED LAW OF THE ITERATED LOGARITHM

R.A. MALLER

Let  $X_i$  be independent and identically distributed random variables with  $S_n = X_1 + X_2 + \dots + X_n$ . We extend a classic result of Kesten, by showing that if  $X_i$  are in the domain of partial attraction of the normal distribution, there are sequences  $\alpha_n$  and  $B(n)$  for which

$$-1 = \liminf_{n \rightarrow +\infty} (S_n - \alpha_n)/B(n) < \limsup_{n \rightarrow +\infty} (S_n - \alpha_n)/B(n) = 1$$

almost surely, and the almost sure limit points of  $(S_n - \alpha_n)/B(n)$  coincide with the interval  $[-1, 1]$ . The norming sequence  $B(n)$  is slightly different to that used by Kesten, and has properties that are less desirable. The converse to the above result is known to be true by results of Heyde and Rogozin.

Let  $X_i$  be independent and identically distributed random variables with distribution  $F$ , and let  $S_n = X_1 + X_2 + \dots + X_n$ . In 1968 Heyde [2] and Rogozin [8] showed that if  $F$  is not in the domain of partial attraction of the normal distribution (cf. Lévy [5], p. 113) then necessarily  $\limsup |S_n - \delta_n|/\gamma(n) = +\infty$  almost surely, or  $(S_n - \delta_n)/\gamma(n) \rightarrow 0$  almost surely, where  $\delta_n$  and  $\gamma(n)$  are constants. Kesten [4] in 1972 proved the converse to this, thus giving the following elegant

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generalisation of the classical law of the iterated logarithm:  $F$  is in the domain of partial attraction of the normal distribution if and only if there is a positive sequence  $\gamma(n)$  and constants  $\delta_n$  satisfying

$$P(S_n \geq \delta_n) \geq \pi, \quad P(S_n \leq \delta_n) \leq \pi \text{ for some } \pi \in (0, 1) \text{ such that}$$

$$-\infty < \liminf_{n \rightarrow +\infty} (S_n - \delta_n) / \gamma(n) < \limsup_{n \rightarrow +\infty} (S_n - \delta_n) / \gamma(n) < +\infty$$

almost surely.

The purpose of the present paper is to give the following extended version of the above result, which partially answers the problem on page 717 of [4].

**THEOREM 1.**  *$F$  is in the domain of partial attraction of the normal distribution if and only if there is a positive sequence  $B(n) \uparrow +\infty$  such that*

$$-1 = \liminf_{n \rightarrow \infty} (S_n - \alpha_n) / B(n) < \limsup_{n \rightarrow \infty} (S_n - \alpha_n) / B(n) = 1$$

almost surely, where  $\alpha_n = n \int_{|u| \leq B(n)} u dF(u)$ . Furthermore the almost sure limit points of  $(S_n - \alpha_n) / B(n)$  are precisely the interval  $[-1, 1]$ .

Our proof of Theorem 1 is based on the methods of Kesten but differs in detail. The norming sequence  $B(n)$  we use, though derived from Kesten's  $\gamma(n)$ , is defined differently and, unlike  $\gamma(n)$ , fails to have the property that  $n^{-\frac{1}{2} + \epsilon} B(n)$  is nondecreasing for  $0 < \epsilon < \frac{1}{2}$ . We discuss this point further following the proof of Theorem 1.

Our method does not depend on symmetrisation of the random variables in an essential way. It can be shown as in [4] that the centering sequence  $\alpha_n$  may be replaced by any sequence  $\delta_n$  for which  $P(S_n \geq \delta_n) \geq \pi$  and  $P(S_n \leq \delta_n) \leq \pi$  for some  $\pi \in (0, 1)$ . Thus  $\alpha_n$  may be replaced by the median of  $S_n$ .

We derive the recurrence part of Theorem 1 from Lemma 1 below, which is a modified version of the criterion of Binmore and Katz (Theorem 2 of [3]) for the recurrence of  $S_n$ . Our result is given in a slightly more general form than is required for the proof of Theorem 1, since it is hoped

that it may have other applications.

We now prove Theorem 1. The statements and proofs of the lemma just mentioned, and of a second lemma we use (which is a version of Lévy's inequality [6, p. 259]), follow this proof.

Proof of Theorem 1. We need only give the sufficiency part of the proof, so suppose  $F$  is in the domain of partial attraction of the normal distribution. Thus there is a sequence  $x_k \uparrow +\infty$  for which

$$\zeta_k = x_k^2 P(|X| > x_k) / V(x_k) \rightarrow 0,$$

where

$$V(x) = \int_{-x}^x u^2 dF(u) - \left[ \int_{-x}^x u dF(u) \right]^2.$$

Fixing  $\varepsilon$  with  $0 < \varepsilon < \frac{1}{4}$  we can assume  $\zeta_k \leq k^{-2/\varepsilon}$ .

Define a sequence  $r_k \uparrow +\infty$  by  $r_k = \left[ \log_2 \left\{ \zeta_k^{-3/4} x_k^2 / V(x_k) \right\} \right]$  ( $[x]$  denotes the integer part of  $x$  and  $\log_2$  the logarithm to base 2), so that

$$2^{r_k} \leq \zeta_k^{-3/4} \left\{ x_k^2 / V(x_k) \right\} = \zeta_k^{1/4} / P(|X| > x_k) < 2 \cdot 2^{r_k},$$

and define  $B(n)$  by

$$B(n) = 2^{r_k/2} \sqrt{2 \log k V(x_k)} \quad \text{when} \quad 2^{r_{k-1}} < n \leq 2^{r_k}.$$

Since  $V$  is ultimately nondecreasing, it is no restriction to assume that  $B(n)$  is nondecreasing for  $n \geq 1$ . Now we truncate at  $x_k$ : let

$X_i^k = X_i$  if  $|X_i| \leq x_k$ , 0 otherwise, let

$$\beta_k = EX_i^k = \int_{-x_k}^{x_k} u dF(u),$$

and note that  $X_i^k$  has variance

$$E \left( X_i^k \right)^2 - \beta_k^2 = \int_{-x_k}^{x_k} u^2 dF(u) - \beta_k^2 = V(x_k) .$$

Hence if  $S_n^k = X_1^k + X_2^k + \dots + X_n^k$ , we have by the Berry-Esseen theorem

(Feller [1, p. 542]), if  $r \geq 1$ ,  $k \geq 1$ , and  $\Phi(x) = (2\pi)^{-\frac{1}{2}} \int_{-\infty}^x e^{-\frac{1}{2}u^2} du$ , that

$$(1) \quad \sup_{-\infty < x < +\infty} \left| P \left\{ S_n^k - 2^r \beta_k < x \sqrt{2^r V(x_k)} \right\} - \Phi(x) \right| \leq 3E \left| X_i^k - \beta_k \right|^3 / 2^{r/2} V^{3/2}(x_k) = L_r^k ,$$

say.

Suppose  $r_k - \varepsilon^{-1} \log_2 k < r \leq r_k$ . Then since for some constants  $c_0$ ,  $c > 0$ ,

$$E \left| X_i^k - \beta_k \right|^3 \leq c_0 E \left| X_i^k \right|^3 \leq c_0 x_k \int_{-x_k}^{x_k} u^2 dF(u) \leq c^2 x_k V(x_k) ,$$

we have

$$\begin{aligned} L_r^k &\leq c \left\{ x_k^2 / 2^r V(x_k) \right\}^{\frac{1}{2}} = c 2^{(r_k - r)/2} \left\{ x_k^2 / 2^{r_k} V(x_k) \right\}^{\frac{1}{2}} \leq c \left\{ 2^{r_k - r + 1} \zeta_k^{3/4} \right\}^{\frac{1}{2}} \\ &\leq c \left\{ 2k^{1/\varepsilon} \zeta_k^{3/4} \right\}^{\frac{1}{2}} \leq c 2^{\frac{1}{2}k} k^{-1/4\varepsilon} , \end{aligned}$$

so that, because  $\varepsilon < \frac{1}{4}$ ,

$$\sum_{k \geq 1} \sum_{r_k - \varepsilon^{-1} \log_2 k < r \leq r_k} L_r^k \leq c 2^{\frac{1}{2}} \varepsilon^{-1} \sum_{k \geq 1} k^{-1/4\varepsilon} \log_2 k < +\infty .$$

Also if  $r_k - \varepsilon^{-1} \log_2 k < r \leq r_k$ ,

$$B(2^r) = 2^{r_k/2} \sqrt{2 \log k V(x_k)} \geq \sqrt{2^r 2 \log k V(x_k)} ,$$

so it follows from (1) with  $x = a\sqrt{2 \log k}$ ,  $a > 1$ , that

$$\begin{aligned} \sum_{k \geq 1} \sum_{r_k - \varepsilon^{-1} \log_2 k < r \leq r_k} P\left\{S_{2^r}^k - 2^r \beta_k > aB(2^r)\right\} \\ \leq \sum_{k \geq 1} \sum_{r_k - \varepsilon^{-1} \log_2 k < r \leq r_k} P\left\{S_{2^r}^k - 2^r \beta_k > a\sqrt{2^r 2 \log k V(x_k)}\right\} \\ \leq \sum_{k \geq 1} \sum_{r_k - \varepsilon^{-1} \log_2 k < r \leq r_k} \left\{L_{2^r}^{k+1} - \Phi(a\sqrt{2 \log k})\right\} \\ < +\infty + 2\varepsilon^{-1} \sum_{k \geq 1} k^{-a^2} \log_2 k < +\infty, \end{aligned}$$

using the approximation  $1 - \Phi(x) \leq 2e^{-\frac{1}{2}x^2}$ , valid when  $x \geq 2$ . It is easy to check that

$$P\left\{S_{2^r}^k - 2^r \beta_k > aB(2^r)\right\} \leq P\left\{S_{2^r}^k - 2^r \beta_k > aB(2^r)\right\} + 2^r P(|X| > x_k),$$

while since

$$\begin{aligned} \sum_{k \geq 1} \sum_{r_{k-1} < r \leq r_k} 2^r P(|X| > x_k) &\leq 2 \sum_{k \geq 1} 2^r k P(|X| > x_k) \\ &\leq \sum_{k \geq 1} \zeta_k^{\frac{1}{2}} \leq \sum_{k \geq 1} k^{-1/2\varepsilon} < +\infty \end{aligned}$$

we have

$$(2) \quad \sum_{k \geq 1} \sum_{r_k - \varepsilon^{-1} \log_2 k < r \leq r_k} P\left\{S_{2^r}^k - 2^r \beta_k > aB(2^r)\right\} < +\infty.$$

Suppose  $r_{k-1} < r \leq r_k - \varepsilon^{-1} \log_2 k$ . By Chebychev's inequality, if  $a > 0$ ,

$$\begin{aligned} \sum_{k \geq 1} \sum_{r_{k-1} < r \leq r_k - \varepsilon^{-1} \log_2 k} P\left\{S_{2^r}^k - 2^r \beta_k > aB(2^r)\right\} \\ \leq a^{-2} \sum_{k \geq 1} \sum_{r_{k-1} < r \leq r_k - \varepsilon^{-1} \log_2 k} B^{-2}(2^r) 2^r V(x_k) \\ = a^{-2} \sum_{k \geq 1} \sum_{r_{k-1} < r \leq r_k - \varepsilon^{-1} \log_2 k} 2^{r-r_k} k (\log k)^{-1} \\ \leq 2a^{-2} \sum_{k \geq 1} (\log k)^{-1} k^{1/\varepsilon} < +\infty \end{aligned}$$

because  $B^2(2^r) = 2^r k (2 \log k V(x_k))$  when  $r_{k-1} < r \leq r_k$ . Again since

$$\sum_{k \geq 1} \sum_{r_{k-1} < r \leq r_k} 2^r P(|X| > x_k) < +\infty, \text{ we can ignore the truncation, and}$$

together with (2) the inequality just derived gives

$$(3) \quad \sum_{k \geq 1} \sum_{r_{k-1} < r \leq r_k} P\left\{S_{2^r}^k - 2^r \beta_k > aB(2^r)\right\} < +\infty \text{ when } a > 1.$$

Now we need the following argument: replacing  $X_i$  by  $-X_i$  we see from (3) that

$$\sum_{k \geq 1} \sum_{r_{k-1} < r \leq r_k} P\left\{S_{2^r}^k - 2^r \beta_k < -aB(2^r)\right\} < +\infty \text{ when } a > 1,$$

so

$$\sum_{k \geq 1} \sum_{r_{k-1} < r \leq r_k} P\left\{\left|S_{2^r}^k - 2^r \beta_k\right| > aB(2^r)\right\} < +\infty \text{ when } a > 1.$$

Letting  $S_n^s$  be a symmetrisation of  $S_n$  in the usual way, this means

$$\sum_{n \geq 1} P\left\{\left|S_n^s\right| > 2aB(2^n)\right\} < +\infty, \text{ and hence by Lemma 2 of Kesten [4],}$$

$\limsup_{n \rightarrow +\infty} \left|S_n^s\right|/B(n) \leq 4$  almost surely. Applying Lemma 4 of [4] now gives

$$S_n^s/B(n) \xrightarrow{P} 0, \text{ so } (S_n - \alpha_n)/B(n) \xrightarrow{P} 0 \text{ where } \alpha_n = n \int_{|u| \leq B(n)} u dF(u) \text{ by}$$

Loève [6, p. 290]. Noting that  $2^{r_k} V(x_k)/x_k^2 > \frac{1}{2} \zeta_k^{-3/4} \rightarrow +\infty$  shows that

$B(2^{r_k})/x_k \rightarrow +\infty$ , and so for  $k$  large enough,

$$\begin{aligned} \left| \frac{2^{r_k} \beta_k - \alpha_{2^{r_k}}}{2^{r_k}} \right| / B(2^{r_k}) &= 2^{r_k} \left| \int_{x_k \leq |u| \leq B(2^{r_k})} u dF(u) \right| / B(2^{r_k}) \\ &\leq 2^{r_k} P(|X| > x_k) \leq \zeta_k^{1/2} \rightarrow 0, \end{aligned}$$

which means we can replace  $2^{r_k} \beta_k$  by  $\alpha_{2^{r_k}}$  in (2) and deduce that

$$\sum_{k \geq 1} P \left\{ S_{\frac{2^{r_k} - \alpha_{2^{r_k}}}{2^{r_k}}} > a B(2^{r_k}) \right\} < +\infty \text{ for } a > 1. \text{ Since } (S_n - \alpha_n) / B(n) \xrightarrow{p} 0,$$

we can apply a version of Lévy's inequality (Lemma 2 below) to obtain from this that for some  $k_0 \geq 1$ ,

$$\sum_{k \geq 1} P \left\{ \max_{k_0 \leq j \leq 2^{r_k}} (S_j - \alpha_j) > a B(2^{r_k}) \right\} < +\infty$$

for  $a > 1$ . By the Borel-Cantelli lemma, then,

$$\limsup_{k \rightarrow +\infty} \max_{k_0 \leq j \leq 2^{r_k}} (S_j - \alpha_j) / B(2^{r_k}) \leq 1$$

almost surely. Now, given any  $n \geq 1$ , choose  $k = k(n)$  so that

$2^{r_{k-1}} < n \leq 2^{r_k}$ ; then  $B(n) = B(2^{r_k})$ , and so we obtain half of what we want:

$$\limsup_{n \rightarrow +\infty} (S_n - \alpha_n) / B(n) \leq \limsup_{n \rightarrow +\infty} \max_{k_0 \leq j \leq 2^{r_k}} (S_j - \alpha_j) / B(2^{r_k}) \leq 1$$

almost surely.

Now let  $a < 1$ ,  $\epsilon > 0$ ,  $a + \epsilon < 1$ ,  $a - \epsilon > 0$ . Applying (1) for

$r = r_k$  gives

$$P\left\{S_{2^k}^k - \alpha_{2^k} \in (a-\epsilon, a+\epsilon)\sqrt{2^k \log kV(x_k)}\right\} \geq (2\pi)^{-\frac{1}{2}} \int_I e^{-\frac{1}{2}u^2} du - 2L_{2^k}^k$$

where  $I = (a-\epsilon, a+\epsilon)\sqrt{2 \log k}$ . Since  $\sum_{k \geq 1} \int_I e^{-\frac{1}{2}u^2} du$  diverges for  $a$  and  $\epsilon$  as defined, while  $\sum_{k \geq 1} L_{2^k}^k < +\infty$ , as shown earlier, we obtain (once again ignoring the truncation)

$$\sum_{k \geq 1} P\left\{S_{2^k}^k - \alpha_{2^k} \in (a-\epsilon, a+\epsilon)B(2^k)\right\} = +\infty.$$

An application of Lemma 1 now will give that  $(S_n - \alpha_n)/B(n)$  is recurrent at  $a$  if  $0 < a < 1$ , providing we verify the conditions of the lemma. If  $\mu_1 > 1$  and  $k_0$  is large enough,

$$B\left[2^{\lceil r k \mu_1 \rceil} / B(2^{r k}) = B(2^{r k + 1}) / B(2^{r k}) \geq \epsilon^{-2}$$

for  $k \geq k_0$ , since it is clearly no loss of generality to assume

$2^{r k + 1} - 2^{r k} \rightarrow +\infty$ . Thus  $B[2^{\lceil r k \mu \rceil} / B(2^{r k}) \geq \epsilon^{-2}$  for  $\mu \geq \mu_1, k \geq k_0$ . If

$1/\mu_1 = \mu_0 \leq 1$  and  $k_0$  is so large that  $2^{r k - 1} < [2^{\lceil r k \mu_0 \rceil}] \leq 2^{r k}$  for

$k \geq k_0, B[2^{\lceil r k \mu_0 \rceil} / B(2^{r k}) = 1$ ; thus  $B[2^{\lceil r k \mu \rceil} / B(2^{r k}) = 1$  for

$\mu_0 \leq \mu \leq 1, k \geq k_0$ . Thus taking  $b_+(\mu) = 1, b_-(\mu) \geq \epsilon^{-2}, \lambda_k = 2^{r k},$

$k \geq k_0$ , we deduce from Lemma 1 that  $(S_n - \alpha_n)/B(n) \in (a-\epsilon, a+\epsilon)$  infinitely often with probability 1, so  $(S_n - \alpha_n)/B(n)$  is recurrent at all points of

$(0, 1)$  and hence  $\limsup_{n \rightarrow +\infty} (S_n - \alpha_n)/B(n) \geq 1$  almost surely. We proved

earlier that  $\limsup_{n \rightarrow +\infty} (S_n - \alpha_n)/B(n) \leq 1$  almost surely, so

$\limsup_{n \rightarrow +\infty} (S_n - \alpha_n)/B(n) = 1$  almost surely. Replacing  $X_i$  by  $-X_i$  shows



that  $(S_n - \alpha_n)/B(n)$  is recurrent at all points of  $(-1, 0)$ , and  $\liminf_{n \rightarrow +\infty} (S_n - \alpha_n)/B(n) = -1$  almost surely. The recurrence at 0 follows since  $(S_n - \alpha_n)/B(n) \xrightarrow{P} 0$ . This completes the proof of Theorem 1.

REMARKS. (i) The sequence  $B(n)$  we used in Theorem 1 is defined differently to Kesten's  $\gamma(n)$ , and it has properties which are less desirable. Since it is constant on large subintervals,  $n^{-\frac{1}{2} + \epsilon} B(n)$  is not nondecreasing for  $0 < \epsilon < \frac{1}{2}$ , whereas  $n^{-\frac{1}{2} + \epsilon} \gamma(n)$  is nondecreasing for  $0 < \epsilon < \frac{1}{2}$ . If we try to use  $\gamma(n)$  in the proof of Theorem 1, we can only obtain a partial result. In fact, if  $x_k$  and  $r_k$  are the sequences defined in the proof, define  $\gamma^*(n)$  by

$$\gamma^*(n) = n^{\frac{1}{2} - \epsilon} 2^{\epsilon r} k \sqrt{2 \log k V(x_k)}, \text{ when } 2^{r k - 1} < n \leq 2^{r k},$$

where  $0 < \epsilon < \frac{1}{2}$  ( $\gamma(n)$  has  $(3/2) \log k$ , instead of  $2 \log k$ ). Then by the method of Theorem 1 we can show, under the same assumptions, that  $(S_n - \alpha_n)/\gamma^*(n)$  is recurrent at all points of  $[-1, 1]$ , and thus that  $\limsup_{n \rightarrow +\infty} (S_n - \alpha_n)/\gamma^*(n) \geq 1$  almost surely; but we only obtain  $\limsup_{n \rightarrow +\infty} (S_n - \alpha_n)/\gamma^*(n) \leq 2$  almost surely. It seems a reasonable conjecture that the result of Theorem 1 actually holds when  $B(n)$  is replaced by  $\gamma^*(n)$ .

(ii) If  $(\lambda_j)$  is a sequence of integers we introduce the notation  $(\lambda_i, i \geq j)$  to mean the collection  $\{\lambda_j, \lambda_{j+1}, \dots\}$ , and we define  $P\{S_{\lambda_j} - \alpha_{\lambda_j} \in I_{\lambda_j} \text{ infinitely often}\}$

$$= \lim_{j \rightarrow +\infty} P\{S_n - \alpha_n \in I_n \text{ for some } n \in (\lambda_i, i \geq j)\}$$

where  $I_n$  are any intervals.

Lemma 1 is a modification of Theorem 2 of [3] in which we allow for centering of  $S_n$ , a wide class of norming sequences, and the fact that recurrence can be deduced from behaviour on a subsequence. We only need to consider nondecreasing norming sequences.

LEMMA 1. Suppose  $\alpha_n$  and  $B(n)$  are constants, with  $B(n) > 0$ ,  $B(n) \uparrow +\infty$ , for which (i) of Lemma 2 holds. Let  $(\lambda_j)$  be a sequence of integers satisfying  $\lambda_j/\lambda_i \geq q_{j-i}$ , where  $q_j \rightarrow +\infty$  as  $j \rightarrow +\infty$ , for which  $B[\mu\lambda_j]/B(\lambda_j) \geq 1/b_+(\mu)$  for  $\mu \in (\mu_0, 1]$  and  $B[\mu\lambda_j]/B(\lambda_j) \geq b_-(\mu)$  for  $\mu \geq 1/\mu_0 > 1$ , for some  $\mu_0 \in (0, 1)$  and some real valued functions  $b_+$  and  $b_-$ . Then if  $b > a \geq 0$ , (i) implies (ii) and (ii) implies (iii):

- (i)  $P\{S_{\lambda_j} - \alpha_{\lambda_j} \in (a, b)B(\lambda_j) \text{ infinitely often}\} = 1$ ;
- (ii)  $\sum_{j \geq 0} P\{S_{\lambda_j} - \alpha_{\lambda_j} \in (a, b)B(\lambda_j)\} = +\infty$ ;
- (iii)  $P\{S_n - \alpha_n \in (a - \epsilon - b/b_-(\epsilon^{-1}), bb_+(1 - \epsilon) + \epsilon)B(n) \text{ infinitely often}\} = 1$  for every  $\epsilon \in (0, 1 - \mu_0)$ .

Proof of Lemma 1. It is easy to see that (i) implies (ii), so let (ii) hold. Fix  $\epsilon > 0$ ,  $\epsilon < 1 - \mu_0$ , and let  $I = (a, b)$ ,  $I' = (a - \epsilon - b/b_-(\epsilon^{-1}), bb_+(1 - \epsilon) + \epsilon)$ . If  $s \geq 1$  is an integer, (ii) implies

$$\begin{aligned} \infty &= \sum_{j \geq s} P\{S_{\lambda_j} - \alpha_{\lambda_j} \in IB(\lambda_j)\} \\ &= \sum_{t=0}^{s-1} \sum_{j \geq 1} P\{S_{\lambda_{js+t}} - \alpha_{\lambda_{js+t}} \in IB(\lambda_{js+t})\} \end{aligned}$$

so there is a  $t \in [0, s)$  for which

$$(4) \quad \sum_{j \geq 1} P\{S_{\lambda_{js+t}} - \alpha_{\lambda_{js+t}} \in IB(\lambda_{js+t})\} = +\infty.$$

Fix  $s$  so large that  $q_s - 1 \geq \epsilon^{-1}$ , and define the disjoint sets

$$E_j = \{S_n - \alpha_n \notin IB(n) \text{ for } n \in (\lambda_{(i+1)s+t}, i \geq j), S_{\lambda_{js+t}} - \alpha_{\lambda_{js+t}} \in IB(\lambda_{js+t})\}$$

so that

$$P(E_j) = \int_{u \in I} P\{S_n^{-\alpha_n} \notin IB(n) \text{ for } n \in (\lambda_{(i+1)s+t}, i \geq j) \mid S_{\lambda_{js+t}}^{-\alpha_{\lambda_{js+t}}} = uB(\lambda_{js+t})\} dP\{S_{\lambda_{js+t}}^{-\alpha_{\lambda_{js+t}}} < uB(\lambda_{js+t})\} .$$

By independence and stationarity, the probability in the integrand equals

$$(5) \quad P\{S_{n-\lambda_{js+t}}^{-\alpha_{n-\lambda_{js+t}}} + (\alpha_{n-\lambda_{js+t}}^{-\alpha_n + \alpha_{\lambda_{js+t}}}) \notin IB(n) - uB(\lambda_{js+t}) \text{ for } n \in (\lambda_{(i+1)s+t}, i \geq j)\} .$$

Now when  $n$  is one of the numbers  $\lambda_{(i+1)s+t}$ ,  $i \geq j$ ,

$$\lambda_{js+t} \leq q_s^{-1} \lambda_{(j+1)s+t} \leq \epsilon n, \text{ and}$$

$$B(n - \lambda_{js+t}) = B(n(1 - \lambda_{js+t}/n)) \geq B((1 - \epsilon)n) \geq B(n)/b_+(1 - \epsilon),$$

which means  $bB(n) - uB(\lambda_{js+t}) \leq bb_+(1 - \epsilon)B(n - \lambda_{js+t})$  when  $u \geq a > 0$ .

Similarly,  $n \geq \lambda_{(j+1)s+t}$  means  $n - \lambda_{js+t} \geq (q_s - 1)\lambda_{js+t} \geq \epsilon^{-1}\lambda_{js+t}$ , so

$$B(n - \lambda_{js+t}) \geq B(\epsilon^{-1}\lambda_{js+t}) \geq b_-(\epsilon^{-1})B(\lambda_{js+t})$$

which means

$$aB(n) - uB(\lambda_{js+t}) \geq aB(n - \lambda_{js+t}) - bB(\lambda_{js+t}) \geq (a - b/b_-(\epsilon^{-1}))B(n - \lambda_{js+t}) .$$

By (i) of Lemma 2, there are integers  $n_0, j_0$ , for which

$$|\alpha_n^{-\alpha_{\lambda_{js+t}}} - \alpha_{n-\lambda_{js+t}}^{-\alpha_{\lambda_{js+t}}}| \leq \epsilon B(n)/b_+(1 - \epsilon) \leq \epsilon B(n - \lambda_{js+t})$$

if  $n > \lambda_{js+t} > j_0$ ,  $n > n_0$ , and  $n$  is one of the numbers  $\lambda_{(i+1)s+t}$ ,  $i \geq j$ . Thus we see that if  $s$  and  $j$  are larger than some fixed integers, the probability in (5) is

$$\begin{aligned} &\geq P\{S_{n-\lambda_{js+t}}^{-\alpha_{n-\lambda_{js+t}}} \notin I'B(n - \lambda_{js+t}) \text{ for } n \in (\lambda_{(i+1)s+t}, i \geq j)\} \\ &\geq P\{S_{n-\lambda_{js+t}}^{-\alpha_{n-\lambda_{js+t}}} \notin I'B(n - \lambda_{js+t}) \text{ for } n \geq \lambda_{(j+1)s+t}\} \\ &\geq P\{S_n^{-\alpha_n} \notin I'B(n) \text{ for } n \geq \lambda_{(j+1)s+t} - \lambda_{js+t}\} \\ &\geq P\{S_n^{-\alpha_n} \notin I'B(n) \text{ for } n \geq q_s - 1\} . \end{aligned}$$

This means

$$\begin{aligned}
 P(E_j) &= P\{S_n - \alpha_n \notin I'B(n) \text{ for } n \geq q_s - 1\} \int_{u \in I} dP\{S_{\lambda_{js+t}} - \alpha_{\lambda_{js+t}} < uB(\lambda_{js+t})\} \\
 &= P\{S_n - \alpha_n \notin I'B(n) \text{ for } n \geq q_s - 1\} P\{S_{\lambda_{js+t}} - \alpha_{\lambda_{js+t}} \in IB(\lambda_{js+t})\}
 \end{aligned}$$

and since the  $E_j$  are disjoint, summing over  $j$  gives

$$1 \geq P\{S_n - \alpha_n \notin I'B(n) \text{ for } n \geq q_s - 1\} \sum_{j \geq 1} P\{S_{\lambda_{js+t}} - \alpha_{\lambda_{js+t}} \in IB(\lambda_{js+t})\}$$

which implies, by (4), that  $P\{S_n - \alpha_n \notin I'B(n) \text{ for } n \geq q_s - 1\} = 0$ . Since  $q_s \rightarrow +\infty$  as  $s \rightarrow +\infty$ , we now have  $P\{S_n - \alpha_n \in I'B(n) \text{ infinitely often}\} = 1$ , as required.

REMARKS. (i) Lemma 1 simplifies if, for example, the  $B(n)$  are assumed to be regularly varying with positive index, that is,  $B(n\mu)/B(n) \rightarrow \mu^\beta$  ( $n \rightarrow +\infty$ ) for  $\mu > 0$  for some  $\beta > 0$ . Then  $\lambda_j$  can be taken to be the geometric subsequence  $[\lambda^j]$ , where  $\lambda$  is any number greater than 1, and (iii) simplifies by omitting  $b_+$  and  $b_-$  altogether (equivalently, putting  $b_-(\epsilon^{-1}) = +\infty$ ,  $b_+(1-\epsilon) = 1$ , which can be achieved at the expense only of replacing  $\epsilon$  by  $2\epsilon$  in (iii)).

(ii) The restriction  $a \geq 0$  can be easily removed from Lemma 1, but it does not seem that (ii) can be replaced by the slightly weaker condition

$$(ii') \sum_{j \geq 0} P\{S_n - \alpha_n \in (a, b)B(n) \text{ for some } n \in [\lambda_j, \lambda_{j+1}]\} = +\infty$$

unless the sequence  $\lambda_j$  grows at most as rapidly as a geometric sequence.

LEMMA 2. Suppose  $(S_n - \alpha_n)/B(n) \xrightarrow{P} 0$ , where  $\alpha_n$  and  $B(n)$  are constants with  $B(n) > 0$ ,  $B(n) \uparrow +\infty$ . Then for every  $\epsilon > 0$ ,  $\epsilon < 1/6$ , there are constants  $n_0(\epsilon)$ ,  $k_0(\epsilon)$ ,  $n_0 > k_0$ , for which  $n \geq n_0$  implies for every real  $x$ ,

$$(i) \max_{k_0 \leq k < n} |\alpha_n - \alpha_k - \alpha_{n-k}| \leq 4\epsilon B(n), \text{ and}$$

$$(ii) (1-\epsilon)P\{\max_{k_0 \leq k \leq n} (S_k - \alpha_k) \geq xB(n)\} \leq P\{S_n - \alpha_n \geq (x-6\epsilon)B(n)\}.$$

Proof of Lemma 2. Let  $0 < \epsilon < 1/6$  and consider

$$\begin{aligned}
 P\{\alpha_n - \alpha_k - \alpha_{n-k} > 4\epsilon B(n)\} &= P\{(S_n - \alpha_n) - (S_k - \alpha_k) - (S_{n-k} - \alpha_{n-k}) - S_n + S_k + S_{n-k} < -4\epsilon B(n)\} \\
 &\leq P\{S_n - \alpha_n < -\epsilon B(n)\} + P\{S_k - \alpha_k > \epsilon B(n)\} \\
 &\quad + P\{S_{n-k} - \alpha_{n-k} > \epsilon B(n)\} + P\{T_{nk} > \epsilon B(n)\}
 \end{aligned}$$

where  $T_{nk} = S_n - S_k - S_{n-k}$ . Since  $(S_n - \alpha_n)/B(n) \xrightarrow{p} 0$  there is a  $k_0(\epsilon) \geq 1$  for which  $P\{|S_k - \alpha_k| > \frac{1}{2}\epsilon B(k)\} < \epsilon$  if  $k \geq k_0$ . Hence when  $n \geq k \geq k_0$ , the first two probabilities in the last expression are each less than  $\epsilon$ . The third probability is also, if  $n - k \geq k_0$ , less than  $\epsilon$ , while if  $n - k \leq k_0$ , it can be made less than  $\epsilon$  by taking  $n \geq n_0$  for some  $n_0(\epsilon) \geq 1$  (whatever the value of  $k \leq n$ ) since only a finite number of the  $X_i$  are being summed. Finally by stationarity

$$P\{T_{nk} > \epsilon B(n)\} = P\left\{\sum_{i=n-k+1}^n X_i - \sum_{i=1}^k X_i > \epsilon B(n)\right\} \leq 2P\{|S_k - \alpha_k| > \frac{1}{2}\epsilon B(n)\} < 2\epsilon$$

if  $n > k \geq k_0$ . Thus we have shown that  $P\{\alpha_n - \alpha_k - \alpha_{n-k} > 4\epsilon B(n)\} < 6\epsilon < 1$  for  $k_0 \leq k < n$ ,  $n \geq n_0$ , so this probability is actually zero. A symmetrical argument gives similarly  $\alpha_n - \alpha_k - \alpha_{n-k} \geq -4\epsilon B(n)$ , and these two together prove (i).

To prove (ii) we proceed as in the proof of Lévy's inequality. Using (i), we have for  $n \geq n_0$ ,

$$\begin{aligned}
 &P\{S_n - \alpha_n \geq (x - 6\epsilon)B(n)\} \\
 &\geq P \bigcup_{j=k_0}^n \left\{ \max_{k_0 \leq i \leq j-1} (S_i - \alpha_i) < xB(n), S_j - \alpha_j \geq xB(n) \right\} \cap \{S_n - S_j - \alpha_{n-j} \geq -2\epsilon B(n)\} \\
 &= \sum_{j=k_0}^n P(A_j)P(B_j), \text{ say.}
 \end{aligned}$$

Note that

$$P(B_j) = P\{S_{n-j} - \alpha_{n-j} \geq -2\epsilon B(n)\} \geq P\{S_{n-j} - \alpha_{n-j} \geq -2\epsilon B(n-j)\} \geq 1 - \epsilon$$

if  $n - j \geq k_0$ , while if  $j \leq n \leq k_0 + j$ ,  $P\{S_{n-j} - \alpha_{n-j} \geq -2\epsilon B(n)\} \geq 1 - \epsilon$  for  $n \geq n_0$  since only a finite number of the  $X_i$  are being summed.

Hence

$$P\{S_n - \alpha_n \geq (x - 6\epsilon)B(n)\} \geq (1 - \epsilon) \sum_{j=k_0}^n P\{A_j\} = (1 - \epsilon)P\{\max_{k_0 \leq k \leq n} (S_k - \alpha_k) \geq xB(n)\},$$

completing the proof. (See [7] for another version of Lévy's inequality.)

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Division of Mathematics and Statistics,  
CSIRO,  
Private Bag,  
PO Wembley,  
Western Australia 6014, Australia.