

Original Article

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Context and domain matter: the error-related negativity in peer presence predicts fear of negative evaluation, not global social anxiety, in adolescents

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Abstract

Background. Social anxiety symptoms are most likely to emerge during adolescence, a developmental window marked by heightened concern over peer evaluation. However, the neurocognitive mechanism(s) underlying adolescent social anxiety remain unclear. Emerging work points to the error-related negativity (ERN) as a potential neural marker of exaggerated self/error-monitoring in social anxiety, particularly for errors committed in front of peers. However, social anxiety symptoms are marked by heterogeneity and it remains unclear exactly what domain(s) of social anxiety symptoms are associated with ERN variation in peer presence, particularly within the adolescent period.

Methods. To advance and deepen the mechanistic understanding of the ERN's putative role as a neural marker for social anxiety in adolescence, we leveraged a social manipulation procedure and assessed a developmentally salient domain of social anxiety during adolescence – fear of negative evaluation (FNE). Adolescents residing in Hanzhong, a small city in the southwestern region of mainland China, had EEG recorded while performing a flanker task, twice (peer presence/absence); FNE, as well as global social anxiety symptoms, was assessed.

Results. Overall ERN increases in peer presence. FNE specifically, but not global levels of social anxiety symptoms, predicted ERN in peer presence.

Conclusions. These data are the first demonstration that the ERN relates to a specific domain of social anxiety in adolescents, as well as the first evidence of such relations within a non-WEIRD (Western, Educated, Industrialized, Rich and Democratic) sample. Results have important implications for theory and research into adolescent social anxiety.

Social anxiety refers to a cluster of symptoms, including an intense fear of social situations or social scrutiny (Morrison & Heimberg, 2013; Rapee & Heimberg, 1997). These symptoms are most likely to emerge during adolescence (Kessler et al., 2005), a developmental window marked by increased social motivation and heightened concern over peer evaluation (Crone & Dahl, 2012; Dahl, 2004; Parker, Rubin, Erath, Wojslawowicz, & Buskirk, 2015; Steinberg & Morris, 2001). Cognitive models suggest that social anxiety involves negative biases in cognitive processing, including biases toward internal sources of social threat (committing errors/mistakes, autonomic responses), and greater self-monitoring within social settings, among other cognitive processes (Clark & Wells, 1995; Pineles & Mineka, 2005; Rapee & Heimberg, 1997). To extend cognitive models and advance a mechanistic understanding of social anxiety, it is important to identify neurocognitive processes linked to social anxiety symptoms in adolescence. In the current study, we focus on the error-related negativity (ERN) as a theoretically relevant neurocognitive process.

Prior work demonstrates that the ERN, an event-related potential (ERP) elicited by errors of commission (Falkenstein, Hohnsbein, Hoormann, & Blanke, 1991; Gehring, Goss, Coles, Meyer, & Donchin, 1993), can serve as a neural marker of exaggerated error monitoring (self-monitoring) in individuals with general anxiety (Hajcak, Klawohn, & Meyer, 2019; Meyer, 2022; Moser, Moran, Schroder, Donnellan, & Yeung, 2013; Olvet & Hajcak, 2008) and social anxiety (Endrass, Riesel, Kathmann, & Buhmann, 2014; Kujawa et al., 2016; Meyer, Carlton, Crisler, & Kallen, 2018). Moreover, the ERN predicts a pattern of error-related autonomic responses that are typically observed in response to threat (Hajcak & Foti, 2008; Hajcak, McDonald, & Simons, 2003). In addition to the ERN, the correct related negativity (CRN) is triggered by correct responses with a similar latency in the response-locked waveform as the ERN but with a smaller amplitude (Ford, 1999; Vidal, Hasbroucq, Grapperon, & Bonnet, 2000, 2003). The CRN appears to have similar morphological and topographical

properties to ERN, and is thought to reflect self-monitoring on correct trials (Allain, Carbone, Falkenstein, Burle, & Vidal, 2004; Suchan, Jokisch, Skotara, & Daum, 2007; Vidal et al., 2000).

Emerging work investigates associations between social anxiety and ERN specifically when errors are committed in social (*v.* non-social) settings (Barker, Troller-Renfree, Pine, & Fox, 2015, 2015; Buzzell et al., 2017b; Voegler et al., 2018). Yet such work has yielded mixed results in terms of the link between global social anxiety symptoms and the ERN, and only one study has investigated these relations during adolescence (Buzzell et al., 2017b). Given that social anxiety symptomology is marked by considerable heterogeneity (Hofmann, Heinrichs, & Moscovitch, 2004; Spokas & Cardaciotto, 2014; Yu, Zhou, Wang, & Tang, 2020), inconsistent associations between the ERN and social anxiety could arise from prior studies focusing on global measures of social anxiety symptoms that collapse across symptom domains. Fear of negative evaluation (FNE) – which refers to apprehension and distress arising from concerns about being judged negatively by others – is thought to reflect a core, defining symptom of social anxiety (Brown & Larson, 2009); Crone & Dahl, 2012; Dahl, 2004; Gerada, 2020; Nelemans et al., 2019; Parker et al., 2015; Steinberg & Morris, 2001; Westenberg, Gullone, Bokhorst, Heyne, & King, 2007), which may more closely relate to ERN variation in social settings. However, the broader cluster of symptoms associated with social anxiety also includes social inhibition/avoidance, feelings of distress within social situations, and less confidence regarding social relationships (Hofmann et al., 2004; Spokas & Cardaciotto, 2014; Yu et al., 2020), as well as potential overlap in symptom presentations from generalized anxiety disorder (Showraki, Showraki, & Brown, 2020). Thus, if the link between the ERN and social anxiety is primarily driven by FNE, then composite measures that collapse across FNE and other symptom domains could obscure such a link, leading to inconsistent findings across studies/samples.

Drawing on cognitive models of social anxiety and prior work investigating the ERN, we propose that ERN variation in social settings is more closely associated with FNE, as opposed to social anxiety symptoms more generally. Cognitive models propose a ‘vicious cycle’ in which FNE leads to greater self-focus, self-monitoring and attentional biases toward social threats (*i.e.* errors/mistakes) within social settings, further exacerbating FNE as the result of a confirmation bias (Clark & Wells, 1995; Rapee & Heimberg, 1997). Thus, cognitive models explicitly predict a close association between FNE and self-monitoring for errors in social settings. Given that the ERN is an established index of self-monitoring for errors, we similarly hypothesize a close association between ERN magnitude within social settings and FNE (as opposed to social anxiety symptoms more generally). Indeed, neural responses associated with anticipated social feedback are correlated with FNE (Topel et al., 2021; Van der Molen et al., 2014), although prior work has not directly assessed the link between FNE and the ERN in social settings.

Adolescence is marked by increased social motivation, increased salience of peer relations (Brown & Larson, 2009), and heightened concern over peer evaluation (Parker et al., 2015; Steinberg & Morris, 2001). FNE not only reflects a hallmark symptom of social anxiety in cognitive models, but also exhibits a normative peak during adolescence (Brown & Larson, 2009); Crone & Dahl, 2012; Dahl, 2004; Gerada, 2020; Nelemans et al., 2019; Parker et al., 2015; Steinberg & Morris, 2001; Westenberg et al., 2007). Thus, our focus here on FNE reflects a theoretically relevant and developmentally salient domain of social anxiety in adolescents.

The primary goal of the current study was to extend neurocognitive understandings of social anxiety in adolescence, as the first study to examine relations among the FNE subdomain of social anxiety symptoms and the ERN in peer presence. In the current study, a sample of adolescents (mean age: 16.98 years, *s.d.* = 0.45) residing in mainland China were assessed for social anxiety symptoms. EEG was recorded while performing a modified flanker task, twice, once while observed by a peer, and once while alone. Our hypotheses were twofold: (1) given normative increases in peer importance during adolescence, participants would exhibit an overall increase in ERN magnitude in peer presence; (2) higher FNE would predict ERN magnitudes in peer presence, whereas global social anxiety symptoms would exhibit diminished or non-significant associations with ERN in peer presence.

Method

Participants and procedure

All participants in this study were recruited from grade 10 at Nanzheng High School, Hanzhong, mainland China. Exclusion criteria were (1) uncorrected visual impairments, (2) a history of head injury with loss of consciousness, and (3) neurological and developmental disorders. All parents and their adolescent children provided written informed consent/assent, and the research procedures were approved by the Shaanxi Normal University Human Research Ethics Board.

The current sample consisted of 30 adolescents (14 male, 16 female; mean age = 16.98 years, *s.d.* = 0.45, range 15.68–17.91). One participant was excluded as the result of feeling unwell on the day of the assessment; two participants due to technical errors in data collection. In the final analyses, 27 participants (13 male, 14 female) were included in behavioral analyses, and 23 participants (10 male, 13 female) were included in EEG analyses.

Families that agreed to participate were invited to the lab. Upon arrival at the lab site, parents and their adolescent children were provided parental consent and child assent. Then parents were brought to another room to complete a demographic survey and the adolescents were asked to complete a modified flanker task (Eriksen & Eriksen, 1974) twice, once while they were being observed by a gender-matched peer who was actually the confederate (the social condition), and once while not being observed (the nonsocial condition). The order of the social and nonsocial conditions was counterbalanced across participants.

In the social condition, the child was informed that there was another participant who would conduct the experiment next to them and wanted to observe their performance to learn more about the task. The ages of the female and male confederates were 16.7 and 18.3 years old, respectively. During the social condition, the confederate was brought to the experiment room and introduced to the participant. To increase the ecological validity of the manipulation of social context and create a typical anxiety-provoking social situation, both confederates were trained to greet participants and make eye contact, but not to initiate any further oral communications, and only provide the simplest answers (*i.e.* yes, no, etc.) while being asked questions by participants. At the end of each block, the confederate read aloud the text feedback presented on the screen without any further comment.

In both conditions, steps were taken to minimize possible social observation effects of the experimenter being in the room while participants completed the flanker task, consistent with

prior work (Barker et al., 2015; Voegler et al., 2018). The experimenter informed the participant that they would not observe their performance and was only there in case of unexpected technical problems. The experimenter then sat 4m away from the participant and pretended to read a book while the participant completed the flanker task.

Task, EEG acquisition, and preprocessing

Flanker task

A modified version of the flanker task (Eriksen & Eriksen, 1974) was administered using e-prime software (Psychology Software Tools, Inc., Sharpsburg, PA, USA). There were eight blocks of 40 trials with a total of 320 trials. The target stimulus was a central arrowhead that pointed to the left or right and was flanked by four additional arrowheads (two on either side of the central arrowhead). Within each block, 50% of trials (20) consisted of target/flanker arrowheads pointing in the same direction (congruent) and 50% trials (20) consisted of arrowheads pointing in the opposite direction (incongruent); congruent and incongruent trials were presented in random order within each block. Arrowheads were presented in Lucida Console, size 240 font. All stimuli were presented for 200 ms followed by a 500 ms response window which ended upon the response. The response–stimulus interval jittered between 1000 and 1300 ms following the response or after 700 ms from stimulus onset. **Figure 1** shows examples of stimulus sequences.

To ensure an adequate number of error trials for response-locked EEG analyses, block-level text-based feedback was provided at the end of each trial block contingent on the participant's accuracy performance (Gehring, Liu, Orr, & Carp, 2012). If the accuracy for that block was below 75%, a message of 'Please improve your accuracy!' was presented; if the accuracy was above 90%, a message of 'Please respond a bit faster' was presented; if the accuracy was in-between 75 and 90%, a message of 'Good job!' was presented. Block-level feedback was presented onscreen the same way within both the nonsocial and social conditions; however, the feedback text was additionally read aloud by the confederate in the social condition.

EEG acquisition

Brain Vision Recorder was used to record 64-channel EEG data, sampled at 500 Hz and acquired using a BrainVision ActiCHamp Amplifier, a high-impedance EEG system designed to maintain high signal-to-noise ratio at relatively high impedance levels. Before data collection, impedances were reduced to below 30 k Ω , consistent with manufacturer recommendations of 25–60 k Ω (Brain Products GmbH, 2016) and prior work (Laszlo, Ruiz-Blondet, Khalifian, Chu, & Jin, 2014). Data were online-referenced to the Cz electrode (note that data were rereferenced to the average of all electrodes as part of later preprocessing to recover EEG variation at/near Cz; Luck, 2014).

EEG preprocessing

EEG data were preprocessed in MATLAB (Mathworks, Natick, MA, USA) using the Maryland Analysis of Developmental EEG (MADE) Pipeline (Debnath et al., 2020). A detailed description of the preprocessing steps employed can be found in the original MADE publication (Debnath et al., 2020). Briefly, MADE consists of a set of MATLAB scripts and EEGLAB (Delorme & Makeig, 2004) functions/plugins, including FASTER (Nolan, Whelan, & Reilly, 2010) and Adjusted-ADJUST (Leach et al., 2020); MADE

performs standard EEG preprocessing steps, including artifact removal via independent component analysis for artifact removal (Winkler, Debener, Müller, & Tangermann, 2015).

Following preprocessing via MADE, EEG data were epoched into 3000 ms segments starting 1000 ms before the response and baseline-corrected to the -400 to -200 ms window. Epochs that exceeded voltages $\pm 100 \mu\text{V}$, if the epochs were recorded from ocular channels (Fp1, Fp2, AF7, and AF8), they were further removed, and if the epochs were recorded from non-ocular channels, they were interpolated at the epoch level. All removed channels were interpolated using the spherical spline interpolation and then the data were referenced to the average of all electrodes. Epochs associated with anticipatory responses (RTs < 200 ms) were removed. Following preprocessing, two participants were removed from all further EEG analyses, as the result of having no usable data in one or more conditions.

Assessment of ERP reliability

To set a minimum threshold for the number of trials required for inclusion in further analyses, ERN reliability in the current dataset was quantified – as a function of trial counts – using the ERP Reliability Analysis (ERA) toolbox (Clayson & Miller, 2017). These analyses demonstrated that a minimum cut-off of seven error trials were needed to achieve a minimum internal consistency of ≥ 0.6 . Applying this trial count threshold, two additional participants were removed from all further EEG analyses, as the result of having <7 trials in one or more conditions. Altogether, 23 participants were included in the final EEG analyses. The mean numbers of trials-by-condition included in subsequent EEG analyses were as follows: 107.83 nonsocial-correct (s.d. = 22.07; range = 43–148), 28.52 nonsocial-error (s.d. = 11.67; range = 7–49), 108.65 social-correct (s.d. = 22.21; range = 58–146) and 29.83 social-error (s.d. = 12.54; range = 8–55).

Measures

Global social anxiety symptom severity

Self-reports of global social anxiety symptoms were assessed via the Social Phobia subscale of a Chinese translation of Screen for Child Anxiety Related Emotional Disorders (SCARED) (Birmaher et al., 1999; Su, Wang, Fan, Su, & Gao, 2008). Note that use of the term 'Social Phobia' is consistent with the original manuscript introducing/validating the SCARED, designed to assess anxiety symptoms in line with the Diagnostic and Statistical Manual (DSM) IV (American Psychiatric Association, 1994; Birmaher et al., 1999). The Social Phobia subscale consists of seven items (e.g. 'It is hard for me to talk with people I don't know well') assessing global social anxiety symptom severity. Each item is rated on a three-point Likert scale, ranging from 0 (not true) to 2 (very true). Items are summed so that higher scores indicate higher levels of social anxiety symptom severity [total possible range: 0–14; scores ≥ 8 reflect the established clinical cut-off for social phobia (social anxiety) disorder (Birmaher et al., 1999)]. Prior work on the SCARED Social Phobia subscale demonstrates moderate-to-high internal consistency and test-retest reliability, and good discriminant validity within a Chinese population (Su, Wang, Fan, Su, & Gao, 2008). Normative data collected on the Social Phobia subscale, within a large sample of Chinese adolescents, are as follows: $M = 4.10$, $s.d. = 3.22$ (Su et al., 2008). The internal consistency of the Social Phobia subscale in the current sample was $\alpha = 0.85$.

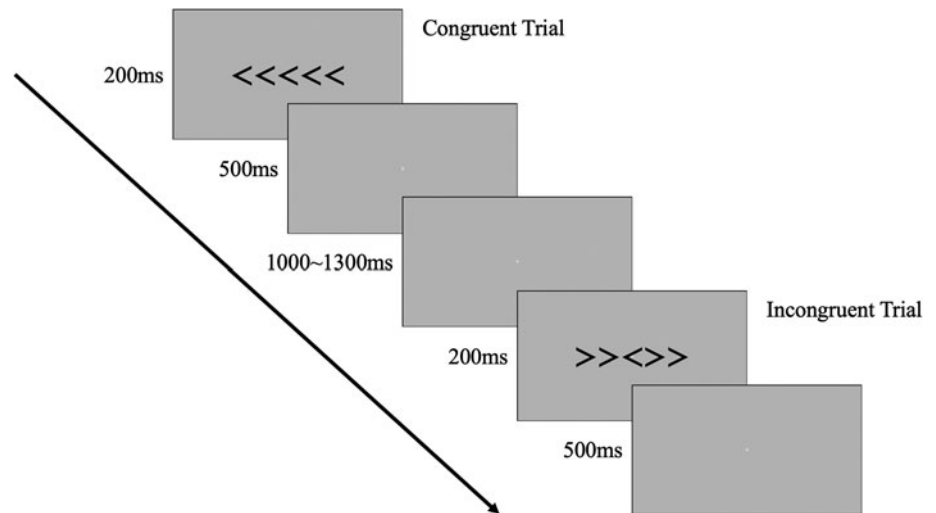


Figure 1. Schematic outline of Flanker task paradigm.

Fear of negative evaluation by peers

Self-reports of FNE were assessed via the FNE subscale of a Chinese translation of the Social Anxiety Scale for Adolescents (SAS-A) (La Greca, 1998). The FNE subscale consists of eight items (e.g. 'I worry about what other kids think of me') assessing adolescents' subjective experiences of fears, concerns, or worries regarding peers' negative evaluations. Each item is rated on a five-point Likert scale, ranging from 1 (not at all) to 5 (all the time). Items are summed so that higher scores indicate higher levels of FNE (total possible range: 8–40). The Chinese version of the SAS has demonstrated appropriate internal consistency, test-retest reliability, and construct validity (Zhou, Xu, Inglés, Hidalgo, & La Greca, 2008). Normative data collected on the FNE subscale, within a large sample of Chinese adolescents, are as follows: $M = 22.73$, $S.D. = 5.9$ (Zhou *et al.*, 2008). The internal consistency of the SAS-FNE subscale in the current sample was $\alpha = 0.94$.

Flanker task behavior

Mean accuracy (ACC) and response time (RT) in each condition (nonsocial and social conditions) were calculated. Participants were required to have greater than 70% accuracy to be included in further behavioral and EEG analyses (all participants met this criterion, mean ACC = 87.23% for the nonsocial condition, and mean ACC = 87.04% for the social condition). Trial-level accuracy and RT data were extracted for further statistical analyses; RT data were first log-transformed to address positive skew (Luze, 1986).

ERPs

Only incongruent trials were included and analyzed (incongruent error trials for ERN, and incongruent correct trials for CRN) in order to isolate error-related effects and avoid stimulus-related confounds (errors are more common on incongruent trials in the flanker task). At the frontocentral electrode (FCz), where error-related brain activity is maximally negative (Meyer, Mehra, & Hajcak, 2021), trial-level mean amplitudes for the CRN and ERN were extracted within the 0–100 ms window following correct and incorrect responses, respectively.

Analytic plan

For all behavioral and EEG analyses, we leveraged a mixed-effects framework in R (R Core Team, 2013), given that this approach is capable of providing accurate and unbiased estimates, even in the presence of low or unequal numbers of trials (Heise, Mon, & Bowman, 2022). For prediction of continuous outcome variables (ERPs, RT), linear mixed effects (LME) modeling was employed, using the nlme package (Pinheiro, Bates, DebRoy, Heisterkamp & Van Willigen, 2016); for prediction of binary accuracy data, a generalized linear mixed effects (GLME) modeling employed, using the glmer function in the lme4 package (Bates, Mächler, Bolker, & Walker, 2014). In all analyses, categorical predictors were contrast coded (–1, +1), continuous predictors converted to Z scores, and effects estimated via maximum likelihood. To report p values, the anova function was used; interaction effects were probed via the emmeans package (Lenth, 2019).

Behavioral analyses

Although not a primary focus of the current study, we carried out basic analyses of behavioral data (RT, accuracy) for completeness. Separate mixed-effects models were fit for accuracy and RT as dependent variables. For accuracy, a GLME model was estimated, with trial-level accuracy data as a binary outcome variable, effects of condition (nonsocial *v.* social), congruency (congruent *v.* incongruent), and their interaction were estimated as fixed effects, and participant entered as a random effect. For RT, an LME model was estimated, with RT as a continuous dependent variable, effects of condition (nonsocial *v.* social), congruency (congruent *v.* incongruent), response (correct *v.* error), and their interactions estimated as fixed effects, and participants entered as a random effect.

ERP analyses

To investigate effects of peer presence on ERN amplitude, we fit an LME model with trial-level ERP amplitudes as the outcome variable, effects of condition (nonsocial *v.* social), response (correct *v.* error), and their interaction estimated as fixed effects, and participant entered as a random effect. As described in the results, this analysis revealed a significant response-by-condition interaction, such that the ERP amplitudes on error trials (i.e. the ERN) but not correct trials (CRN) were increased in the social

v. nonsocial context. Thus, subsequent modelling focused on error trials only (ERN).

To examine effects of FNE *v.* global anxiety symptoms on the ERN in peer presence/absence, we fit two separate LME models employing either FNE or global anxiety symptoms as predictors and directly compared relative model fit via Akaike information criterion (AIC) model selection. In each case, an LME model was fit, with trial-level ERN amplitudes as the outcome variable, condition (nonsocial *v.* social) included as a fixed effect, and participant entered as a random effect. The FNE model additionally included FNE, as well as its interaction with condition, as fixed effects; the global social anxiety symptoms model included global social anxiety symptoms, as well as its interaction with condition, as fixed effects.

Given that the FNE model exhibited superior fit (see Results), we analyzed/interpreted results of this model; results of the global social anxiety model are reported in the online Supplementary material, for completeness. We additionally carried out an online Supplemental analysis in which global social anxiety symptoms were added to the FNE model as a control variable.

Results

Self-report questionnaires

The mean SCARED social anxiety score was 7.52 (s.d. = 3.50, range = 1–14), and the mean FNE score was 21.11 (s.d. = 8.74, range = 8–40). Although our analyses focus on continuous measures of social anxiety symptoms, consistent with a dimensional approach, it is worth noting that 12 of the 23 participants included in EEG analyses exceeded the established clinical cut-off for social phobia (social anxiety) disorder on the SCARED-SP (Birmaher et al., 1999). FNE was significantly correlated with SCARED social anxiety ($r = 0.592$, $p < 0.001$).

Behavioral results

The mean [\pm standard deviation (s.d.)] accuracies in the nonsocial and social conditions were $87.23 \pm 4.34\%$ and $87.04 \pm 4.61\%$, respectively (see Table 1). GLME modeling revealed a main effect of congruency, with participants responding less accurately in incongruent trials [$t(16\ 276) = 28.805$, $p < 0.001$]. There was no significant main effect of social condition [$t(16\ 276) = 0.650$, $p = 0.516$], nor a congruency-by-condition interaction [$t(16\ 276) = 1.716$, $p = 0.086$].

Table 1. Descriptive statistics characterizing behavior and ERPs as a function of condition

	Non-social condition (<i>N</i> = 27)	Social condition (<i>N</i> = 27)
Behavior measures		
Accuracy (%)	87.23 (4.34)	87.04 (4.61)
Correct reaction time (ms)	350.02 (35.78)	350.20 (34.08)
Error reaction time (ms)	295.71 (30.50)	294.31 (27.98)
ERPs (μ V)	(<i>N</i> = 23)	(<i>N</i> = 23)
CRN	−0.87 (2.59)	−0.79 (2.48)
ERN	−4.52 (3.08)	−5.39 (3.44)

Note. Values reflect mean scores (standard deviation in parentheses).

Mean RTs for the nonsocial and social conditions were 343.15 ± 36.36 ms and 343.02 ± 34.49 ms, respectively. LME modeling identified a main effect of response, with participants responding faster in error trials [$t(16\ 272) = 37.374$, $p < 0.001$], and a main effect of congruency, with participants responding faster in congruent trials [$t(16\ 272) = -18.249$, $p < 0.001$]. Moreover, there was a response-by-congruency interaction effect [$t(16\ 272) = -4.295$, $p < 0.001$]; within correct trials, participants responded slower in incongruent responses [$t(16\ 272) = 39.365$, $p < 0.001$], within error trials, participants responded slower in incongruent responses [$t(16\ 272) = 7.273$, $p < 0.001$]. No significant main effect of social condition or other interaction effects were identified (all $ps > 0.116$). No correlations between behavioral data (accuracy and RT) and self-report questionnaires were statistically significant (all $ps > 0.223$).

ERP results

Descriptive statistics for CRN and ERN in the nonsocial and social conditions are reported in Table 1. Figure 2 displays the ERP waveforms and scalp topography for the nonsocial and social conditions across participants at the FCz electrode (where the ERN was maximal).

Results of the LME model assessing fixed effects of condition (nonsocial *v.* social), response (correct *v.* error), and their interaction on trial-level ERP (ERN/CRN) amplitudes are reported in Table 2. This model revealed a significant main effect of response [$\beta = 0.288$, 95% CI (0.260–0.317), s.e. = 0.014, $p < 0.001$], qualified by a response-by-condition interaction [$\beta = 0.037$, 95% CI (0.009–0.065), s.e. = 0.014, $p = 0.009$]; the nature of this interaction was such that the ERN was significantly larger in the social (compared to nonsocial) condition [$t(6295) = 2.393$, $p = 0.017$], whereas the CRN did not differ across conditions [$t(6295) = -1.073$, $p = 0.283$]. Given that social context only exhibited a significant effect for error trials (ERN), and not correct trials (CRN), subsequent modelling focused only on error trials (ERN).

We carried out AIC model selection to compare the following LME models: (1) the FNE model, defined by fixed effects of FNE, condition (nonsocial *v.* social), and their interaction, on trial-level ERN amplitudes; (2) the global social anxiety model, defined by fixed effects of global social anxiety, condition (nonsocial *v.* social), and their interaction, on trial-level ERN amplitudes. AIC model selection revealed that FNE model, carrying 81% of the cumulative model weight, had lower AIC (AIC = 8933.95) than global social anxiety model (AIC = 8936.83), indicating that the FNE model provided a better fit to the data. Given that the FNE model exhibited superior fit, subsequent analyses focused on interpretation of the results of this model (however, for completeness, see the online Supplementary material for results of the global social anxiety model).

Results of the LME model assessing fixed effects of condition (nonsocial *v.* social), FNE, and their interaction on trial-level ERN amplitudes are reported in Table 3. This model revealed a significant main effect of condition [$t(1317) = -2.672$, $p = 0.008$], qualified by a condition-by-FNE interaction [$\beta = -0.052$, 95% CI (−0.102 to −0.002), s.e. = 0.026, $p = 0.042$]; the nature of this interaction was such that higher levels of FNE were more strongly associated with a larger (more negative) ERN in the social (compared to nonsocial) condition [$t(1317) = 2.032$, $p = 0.042$] (See Figure 3). We additionally carried out an online Supplemental analysis in which global social anxiety symptoms were added to

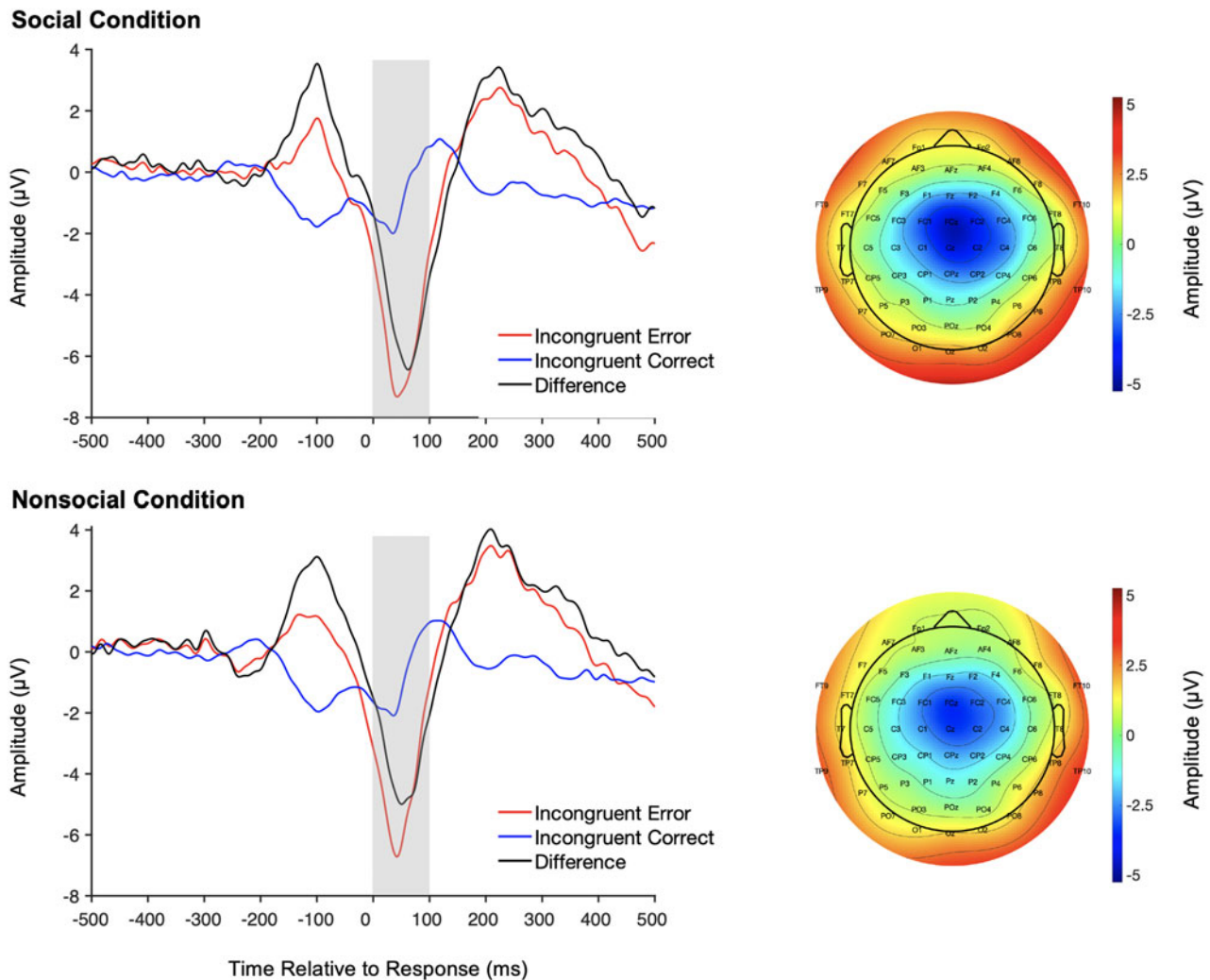


Figure 2. Averaged ERP and topographic plots for the social and nonsocial conditions. ERP plots depict averaged ERPs for incongruent-error trials (ERN: red), incongruent-correct trials (CRN: blue), and their difference (black), as a function of condition (top: social; bottom: nonsocial). Topographic plots corresponding to the shaded region within each ERP plot (0–100 ms) depict mean amplitude differences between incongruent-error and incongruent-correct trials, separately for each condition (top: social; bottom: nonsocial).

the FNE model, described above, as a control variable. As reported in the online Supplementary material, the condition-by-FNE interaction remained significant when controlling for global social anxiety symptoms [$\beta = -0.052$, 95% CI (-0.102 to -0.002), *s.e.* = 0.026, $p = 0.042$], STATA, demonstrating that FNE predicted unique variance in ERN, while controlling for global social anxiety symptoms.

Discussion

Consistent with our hypotheses, we observed an overall increase in ERN magnitude in peer presence. Comparing the statistical models involving either FNE or global social anxiety symptoms to predict peer/alone ERN variation revealed a better fit for the FNE model. Moreover, higher levels of FNE were predictive of an increased ERN in the peer (*v.* alone) setting – an effect that held even when controlling for global social anxiety symptoms. In contrast, global social anxiety symptoms did not significantly relate to the ERN. Results of the current study are consistent with cognitive models of social anxiety (Clark & Wells, 1995;

Rapee & Heimberg, 1997), which propose a close link between FNE and self-monitoring for errors within social settings, providing a crucial neurocognitive extension of prior work. Moreover, these findings are consistent with developmental theory describing an increase in the importance of peers and normatively high levels of FNE during adolescence (Brown & Larson, 2009; Crone & Dahl, 2012; Dahl, 2004; Gerada, 2020; Nelemans *et al.*, 2019; Parker *et al.*, 2015; Steinberg & Morris, 2001; Westenberg *et al.*, 2007). These findings have implications for theory and measurement of adolescent social anxiety, including the importance of considering error sensitivity (Chong & Meyer, 2019) and the possibility of targeting error-related processing (Meyer *et al.*, 2020) in adolescent social anxiety. These data also highlight the ERN's utility as a neurocognitive index of self-monitoring within social settings, which complements self-report measures and could be leveraged by future studies seeking to examine causal mechanisms predicted by cognitive models of social anxiety. We discuss these ideas in further detail below.

Given that social anxiety is characterized by a cluster of heterogeneous symptoms (Inderbitzen-Nolan & Walters, 2000; La Greca

Table 2. Model 1: ERP ~ conditions × response

Predictors	β [95% CI]	s.e.	t	p
(Intercept)	-0.167 [-0.301 to -0.033]	0.068	-2.442	0.015*
Conditions	-0.023 [-0.051 to 0.005]	0.014	-1.626	0.104
Responses	0.288 [0.260–0.317]	0.014	20.130	<0.001***
Conditions × responses	0.037 [0.009–0.065]	0.014	2.620	0.009**

Observations = 6321, $N = 23$.
 Note: * $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$.

Table 3. Model 2: ERN ~ conditions × FNE

Predictors	β [95% CI]	s.e.	t	p
(Intercept)	-0.461 [-0.623 to -0.299]	0.083	-5.580	<0.001***
Conditions	-0.069 [-0.119 to -0.018]	0.026	-2.672	0.008**
FNE	-0.163 [-0.338 to 0.012]	0.084	-1.935	0.067
Conditions × FNE	-0.052 [-0.102 to -0.002]	0.026	-2.032	0.042*

FNE, fear of negative evaluation.
 Observations = 1342, $N = 23$.
 Note: * $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$.

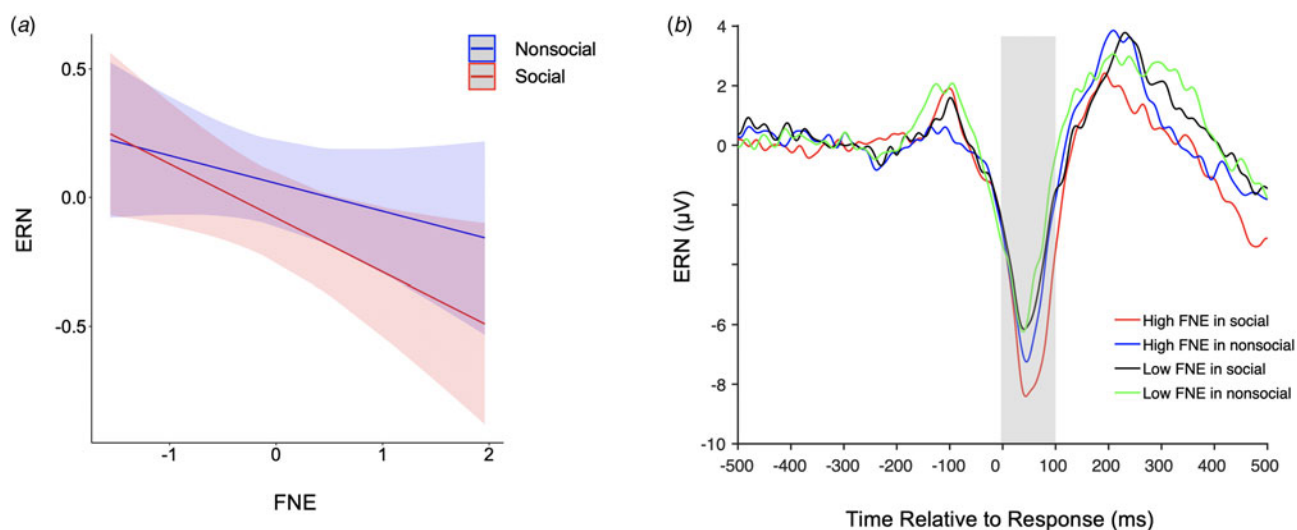


Figure 3. Association between FNE and ERN. (a) Line plot depicting associations between FNE and the ERN, separately for the nonsocial (red) and social (blue) conditions. (b) ERN ERP plots as a function of FNE group and condition (nonsocial v. social). For plotting purposes only, participants were grouped into high and low FNE groups based on the median value of FNE scores (median = 18). ERN plots depict ERN ERPs for: the high FNE group in the social condition (red), high FNE group in the nonsocial condition (blue), low FNE group in the social condition (black), and low FNE group in the nonsocial condition (green). The shaded region reflects the 0–100 ms time window.

& Lopez, 1998; Watson & Friend, 1969), it is reasonable to assume that no single neurocognitive measure would relate to all social anxiety symptoms equally (or at least not reliably across study samples); results of the current study are in line with this notion. Moreover, our data suggest that successful identification of links between symptom domains and neurocognitive measures requires precision at the level of symptom assessment as well as taking into account the situational context of neural assessment and the broader developmental context. Toward these ends, the current study focused on relations between a developmentally salient symptom domain of social anxiety – FNE (e.g. La Greca

& Lopez, 1998; Watson & Friend, 1969) and neural activity associated with error-monitoring (self-monitoring) under peer observation (e.g. Hajcak, Moser, Yeung, & Simons, 2005; Kim, Iwaki, Uno, & Fujita, 2005), within the context of adolescence – when FNE and the importance of peers reach peak levels (Brown & Larson, 2009); Crone & Dahl, 2012; Dahl, 2004; Gerada, 2020; Nelemans et al., 2019; Parker et al., 2015; Steinberg & Morris, 2001; Westenberg et al., 2007).

Importantly, the current study demonstrates that an increased ERN in social settings is associated with the FNE symptom domain in adolescents. Yet, we do not suggest that the ERN

reflects the neurocognitive basis of FNE *per se*. Prior work demonstrates that the ERN indexes a neurocognitive process underlying error-monitoring (self-monitoring) (Agam et al., 2011; Buzzell et al., 2017a; Dehaene, Posner, & Tucker, 1994; Falkenstein et al., 1991; Gehring et al., 1993; Miltner, Braun, & Coles, 1997; Ullsperger, Fischer, Nigbur, & Endrass, 2014). Similarly, we suggest that an increased ERN in social settings indexes increased self-monitoring. As noted, cognitive models propose a vicious cycle whereby greater FNE leads to greater self-monitoring within social settings, which in turn further strengthens/worsens FNE (Clark & Wells, 1995; Rapee & Heimberg, 1997). Thus, we argue that the link between FNE and the ERN in peer presence reflects a neurocognitive extension of cognitive models of social anxiety. These models further state that social anxiety is specifically associated with increased self-monitoring for negative aspects of one's behavior (e.g. errors) as opposed to a general increase in self-monitoring (Clark & Wells, 1995; Rapee & Heimberg, 1997). Our data provide converging neural evidence to support this supposition, as peer presence and FNE predicted larger increases in the ERN (self-monitoring on error trials), as opposed to the CRN (self-monitoring on correct trials). As noted by cognitive models (Clark & Wells, 1995; Rapee & Heimberg, 1997), increased self-monitoring of negative aspects of one's behavior could explain why social anxiety is maintained/worsened across repeated social encounters in daily life: social encounters are experienced through the lens of one's (neuro-) cognitive biases. A bias toward greater self-monitoring for errors in social settings may serve to confirm one's initial fears of negative evaluation, strengthening/worsening FNE.

This is the first study to directly investigate relations between FNE and the ERN in peer presence among adolescents. Given our cross-sectional study design, it remains unknown whether the association between FNE and ERN remains stable across development. Emerging developmental theory on the nature of anxiety-ERN relations argues that whereas the ERN-anxiety link in childhood is primarily driven by external threats/fears, ERN-anxiety relations fundamentally change with the transition to adolescence – to self-conscious shyness and worry about behavioral competence and social evaluation (Meyer, 2017, 2022). Meyer et al. have shown that adolescent ERN-anxiety relations are primarily driven by social anxiety symptoms, as opposed to symptoms of other anxiety disorders (Meyer et al., 2018). Moreover, normative ERN increases during the transition to adolescence are mediated by (normative) increases in subclinical social anxiety symptoms (Meyer et al., 2018). Based on the results of the current study, taken together with the normative increase in FNE during the transition to adolescence (Gerada, 2020; Nelemans et al., 2019; Westenberg et al., 2007), we propose that ERN variation becomes more closely coupled with individual differences in FNE across adolescence. Longitudinal studies that measure ERN and FNE across the peri-adolescent window are needed to evaluate this proposal.

The current study has several implications for clinical neuroscience research and efforts to develop novel assessment approaches. Broadly, the current findings provide additional emphasis on the importance of understanding error sensitivity (Chong & Meyer, 2019) and specifically targeting error-related processing (Meyer et al., 2020) in pediatric anxiety, particularly during adolescence. Moreover, assessment of the ERN in peer presence may reflect a promising neural marker, indicating the degree to which an individual exhibits greater self-monitoring in social settings. As such the ERN could provide a complement

to self-report measures and be particularly useful in settings or participant groups more likely to be impacted by metacognitive awareness, social desirability, or recall biases (Brewin, Andrews, & Gotlib, 1993; Van de Mortel, 2008). While we argue that the ERN underlies self-monitoring, as opposed to a direct index of FNE, cognitive models of social anxiety (Clark & Wells, 1995; Rapee & Heimberg, 1997) and converging neurocognitive evidence from the current study establish a close association between self-monitoring in peer presence and FNE. Thus, the ERN could also be employed as an (indirect) neural marker of FNE levels. Going one step further, an intriguing possibility is that repeated assessments of FNE and ERN, at the intraindividual level, could provide the means to directly capture the 'vicious cycle' proposed by cognitive models of social anxiety (Clark & Wells, 1995; Rapee & Heimberg, 1997). However, further work examining intraindividual relations among FNE and ERN is needed, to include dissociating trait- *v.* state-based associations in these constructs. There is also a need for additional work studying these constructs via methods capable of identifying causal pathways (e.g. longitudinal designs or causal manipulation studies). Such work would provide strong evidence for whether/how FNE and the ERN causally influence one another, and a critical test of whether these constructs are implicated in a 'vicious cycle' in which they mutually strengthen/worsen one another.

Strengths, limitations, and future directions

Strengths of the current study include the targeted assessment of a theoretically relevant and developmentally salient symptom domain of social anxiety – FNE, inclusion of a non-WEIRD (Western, Educated, Industrialized, Rich and Democratic) (Henrich, Heine, & Norenzayan, 2010) sample of adolescents, and at the theoretical level, providing a neurocognitive extension of cognitive models of social anxiety. Yet, this study has several limitations. The first limitation is that our sample size was relatively small, resulting in low statistical power to detect small effects; thus, results of the current report should be interpreted in light of the small sample size. We were also not sufficiently powered to test whether biological sex or gender moderates observed links between ERN, FNE and social context. Similarly, the age range of participants was selected to be narrow in the current study to minimize possible age-related effects, as we would similarly be underpowered to test whether age moderates links between ERN, FNE and social context. Therefore, future work is needed to replicate our findings within a larger sample, and should adopt a broader sampling strategy to investigate whether the reported results are moderated by biological sex, gender, or age. Note that additional limitations of the current study are that, although *a priori*, hypotheses were not pre-registered, and there were no planned analyses of power to determine sample size. Thus, attempts to replicate the current results could further benefit from pre-registering hypotheses, and leverage effect sizes reported in the current study to carry out *a priori* power analyses. It is also worth noting that the current study employed a standard, lab-based assessment of error monitoring via the Flanker task, assessed at a single point in time. Future work could extend this approach via repeated, longitudinal assessments within the lab, or even work to develop recording/analysis protocols that would allow for more continuous, *in vivo* assessment of these constructs outside the laboratory (Caricato et al., 2020). While a strength of the current study is the inclusion of a non-WEIRD sample, it remains unknown whether associations between social

anxiety symptomology and the ERN in peer presence are stable across cultures; cross-cultural studies are needed. Lastly, when interpreting the findings, the cross-sectional design of the current study should be taken into account, since it does not allow us to draw conclusions on the causal direction of associations between FNE and ERN. Regardless, the current study has important implications for clinical theory and social anxiety research; it demonstrates the utility of mapping neurocognitive construct(s) onto specific symptom domain(s) of social anxiety, and highlights the importance of carefully considering precision in symptom assessment, the situational context of neural measurement, and the developmental backdrop in which such work is conducted.

Supplementary material. The supplementary material for this article can be found at <https://doi.org/10.1017/S0033291723000466>

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Ethical standards. The authors assert that all procedures contributing to this work comply with the ethical standards of the relevant national and institutional committees on human experimentation and with the Helsinki Declaration of 1975, as revised in 2008.

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