

ON UPCROSSING PROBABILITIES

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1. Introduction. In [1] a simple but ingenious technique was developed for calculating hitting probabilities for submartingales (or martingales or supermartingales) subject to various constraints. This technique is extended here in order to find sharp bounds on upcrossing probabilities for submartingales subject to constraints. The general results in Section 2 are applied to submartingales $\{X_n\}_{n=1}^\infty$ such that $E[(X_n - a)^+]^p \leq L$ (a constant) for all n , $p \geq 1$ and we find the probability of at least k upcrossings of $[a, b]$ is at most

$$\frac{L - [(m - a)^+]^p}{(b - a)^p} \frac{(p - 1)^{p-1}}{\left(k + \frac{(m - a)^+}{b - a} (p - 1) - 1\right)^p + (p - 1)^{p-1}},$$

where $m = EX_1$. For $p = 1$ this bound collapses to $(L - (m - a)^+) / ((b - a)k)$ (taking $(p - 1)^{p-1} = 1$ when $p = 1$). A simple corollary is that Doob's upcrossing inequality is sharp. A second example gives Dubins' sharp bounds on upcrossing probabilities for bounded martingales.

2. General Results. Keeping the notation established in [1] let R be the set of real numbers; B be the Borel subsets of R ; $R^\infty = R \times R \times \dots$; and $B^\infty = B \times B \times \dots$. Let $\{X_n\}_{n=1}^\infty$ be the coordinate process on R^∞ . A submartingale (or martingale or supermartingale) may be regarded as a probability measure P on B^∞ . $\{X_n\}_{n=1}^\infty$ defined on $\{R^\infty, B^\infty, P\}$ is a submartingale in the usual sense.

Let μ be a probability measure on (R, B) ; we define the following classification:

Definition 1a.) μ satisfies a *condition of type* $(\phi, \underline{r}, \bar{r})$ if there is a family ϕ of convex, increasing, Borel functions from R to R and mappings \underline{r} and \bar{r} of ϕ to $R \cup \{-\infty, \infty\}$ such that for all $\theta \in \phi$,

$$\underline{r}(\theta) \leq \int \theta(x)\mu(dx) \leq \bar{r}(\theta).$$

(ϕ, \bar{r}) means $\underline{r}(\theta) = -\infty$ for all $\theta \in \phi$.

1.b) μ satisfies a *condition of type* (L, U) if there exist two constants $L < U$ such that $\mu\{[L, U]\} = 1$.

Definition 2.a) A probability measure Q on (R^∞, B^∞) satisfies a *condition of*

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type $(\phi, \underline{r}, \bar{r})$ if for all $\theta \in \phi$ and all n ,

$$\underline{r}(\theta) \leq \int \theta(X_n)dQ \leq \bar{r}(\theta).$$

2.b) Q satisfies a condition of type (L, U) if $Q\{L \leq X_n \leq U\} = 1$ for all n .

LEMMA 1. Let M be the collection of all submartingales satisfying conditions of type $(\phi, \underline{r}, \bar{r})$ and/or type (L, U) (that is certain conditions of these types are satisfied). If $\tau_1 \leq \tau_2 \leq \dots \leq \tau_m$ are bounded stopping times and $P \in M$, then the measure Q on B^∞ defined by

$$(1) \quad Q\{X_1 \in B_1, X_2 \in B_2, \dots, X_m \in B_m, X_{m+1} \in B_{m+1}, \dots, X_{m+l} \in B_{m+l}\} \\ P\{X_{\tau_1} \in B_1, X_{\tau_2} \in B_2, \dots, X_{\tau_m} \in B_m, X_{\tau_m} \in B_{m+1}, \dots, X_{\tau_m} \in B_{m+l}\}$$

belongs to M .

Proof. Q is well defined since the stopping times are bounded. Next consider any cylinder set measurable with respect to X_1, X_2, \dots, X_k ; say

$$C = \{X_1 \in B_1, X_2 \in B_2, \dots, X_k \in B_k\}.$$

$$(2) \quad \int_C E_Q\{X_{k+1}|X_k, \dots, X_1\}dQ = \int_C X_{k+1} dQ = \int_{C'} X_{\tau_{k+1}} dP \text{ by (1),}$$

where $C' = \{X_{\tau_1} \in B_1, X_{\tau_2} \in B_2, \dots, X_{\tau_k} \in B_k\}$. Since P is a submartingale,

$$\int_{C'} X_{\tau_{k+1}} dP \geq \int_{C'} X_{\tau_k} dP = \int_C X_k dQ.$$

Since the σ -algebra of sets measurable with respect to X_1, X_2, \dots, X_k ($\sigma(X_1, \dots, X_k)$) is generated by sets of the form C , we have

$$\int_A E_Q\{X_{k+1}|X_k, \dots, X_1\}dQ \geq \int_A X_k dQ,$$

where $A \in \sigma(X_1, \dots, X_k)$. Therefore Q is a submartingale. If $\theta \in \phi$ then $\{\theta(X_n)\}_{n=1}^\infty$ defined on $\{R^\infty, B^\infty, P\}$ is a submartingale. Hence

$$\underline{r}(\theta) \leq \int \theta(X_1)dP \leq \int \theta(X_{\tau_k})dP = \int \theta(X_k)dQ \text{ for all } k.$$

Moreover if l is an integer such that $\tau_m \leq l$ then

$$\int \theta(X_k)dQ = \int \theta(X_{\tau_k})dP \leq \int \theta(X_1)dP \leq \bar{r}(\theta).$$

Therefore Q satisfies condition $(\phi, \underline{r}, \bar{r})$. Condition (L, U) follows trivially. Therefore $Q \in M$. This completes the proof.

Let B_1, B_2, \dots, B_m be Borel sets. Let

$$T = \{X_{n_1} \in B_1 \text{ for some } n_1, X_{n_2} \in B_2 \text{ for some } n_2 \geq n_1, \dots, X_{n_m} \in B_m \text{ for some } n_m \geq n_{m-1}\}.$$

THEOREM 1. *Let M be the collection of all submartingales satisfying certain conditions of type (ϕ, ϱ, \bar{r}) and/or type (L, U) . Then*

$$\sup_{P \in M} P\{T\} = \sup_{P \in M} P\{X_1 \in B_1, \dots, X_m \in B_m\}.$$

Proof. For $\omega = (x_1, x_2, \dots) \in R^\infty$, let

$$\begin{aligned} \tau_1 &= \begin{cases} \text{least } n_1 \text{ (if any) such that } x_{n_1} \in B_1, \\ \infty \text{ if there is no such } n_1; \end{cases} \\ \tau_2 &= \begin{cases} \text{least } n_2 \geq n_1 \text{ (if any) such that } x_{n_2} \in B_2, \\ \infty \text{ if there is no such } n_2; \end{cases} \\ \tau_m &= \begin{cases} \text{least } n_m \geq n_{m-1} \text{ (if any) such that } x_{n_m} \in B_m, \\ \infty \text{ if there is no such } n_m. \end{cases} \end{aligned}$$

Therefore $\tau_1 \leq \tau_2 \leq \dots \leq \tau_m$ and

$$\begin{aligned} P\{T\} &= P\{\tau_m < \infty\} = \lim_{n \rightarrow \infty} P\{\tau_m \leq n\} \\ &= \lim_{n \rightarrow \infty} P\{X_{\tau_1 \wedge n} \in B_1, \dots, X_{\tau_m \wedge n} \in B_m\}. \end{aligned}$$

However, by Lemma 1.a),

$$P\{X_{\tau_1 \wedge n} \in B_1, \dots, X_{\tau_m \wedge n} \in B_m\} = Q\{X_1 \in B_1, \dots, X_m \in B_m\}$$

for some $Q \in M$. Hence

$$\lim_{n \rightarrow \infty} P\{X_{\tau_1 \wedge n} \in B_1, \dots, X_{\tau_m \wedge n} \in B_m\} \leq \sup_{P \in M} P\{X_1 \in B_1, \dots, X_m \in B_m\},$$

and we have

$$\sup_{P \in M} P\{T\} \leq \sup_{P \in M} P\{X_1 \in B_1, \dots, X_m \in B_m\}.$$

The reverse inequality is immediate, completing the proof.

Theorem 1 provides a prescription for obtaining sharp bounds for upcrossing probabilities. Essentially it says stopping times are unnecessary.

For any pair of real numbers $a < b$ and any $\omega = (x_1, x_2, \dots) \in R^\infty$ define γ_{ab} to be the number of upcrossings of the interval $[a, b]$. Define

$$\begin{aligned} S_n &= \{X_1 \leq a, X_2 \geq b, \dots, X_{2n-1} \leq a\} \quad \text{and} \\ T_n &= \{X_1 \leq a, X_2 \geq b, \dots, X_{2n-1} \leq a, X_{2n} \geq b\} \end{aligned}$$

for all n .

LEMMA 2.a. Let a, c, r and W be reals such that $a < c$ and $W \in [0, 1)$; let $l \in R$ be defined by $Wa + (1 - W)l = c$ and let β be the two point probability $\beta = W\delta_a + (1 - W)\delta_l$ (δ_a and δ_l are point probabilities at a and l respectively). Then among all probabilities μ on R such that $\mu((-\infty, a]) \geq W$ and $\int x\mu(dx) \geq c$, β minimizes

$$\theta(r) \cdot \mu(-\infty, r] + \int_{x>r} \theta(x) \mu(dx) = \int \theta(r) \vee \theta(x) \mu(dx)$$

whatever convex, increasing function θ may be.

Proof. Let \mathcal{C} be the class of convex, increasing polygonal functions with a finite number of vertices. It is clear that for any convex increasing function θ ,

$$\int \theta(r) \vee \theta(dx) = \sup_{\substack{g \in \mathcal{C} \\ g \leq \theta}} \int \theta(r) \vee g(x) \mu(dx).$$

Thus to show $\int \theta(r) \vee \theta(x) \mu(dx) \geq \int \theta(r) \vee \theta(x) \beta(dx)$ it suffices to show $\int \theta(r) \vee g(x) \mu(dx) \geq \int \theta(r) \vee g(x) \beta(dx)$ for all $g \in \mathcal{C}$. Now if $g \in \mathcal{C}$, then $\theta(r) \vee g$ may be represented in the following form:

$$\begin{aligned} \theta(r) \vee g(x) = \theta(r) + d_1(x \vee r - r) + (d_2 - d_1)(x \vee x_1 - x_1) + \dots \\ + (d_n - d_{n-1})(x \vee x_n - x_n), \end{aligned}$$

where $\{(r, \theta(r)), (x_1, g(x_1)), \dots, (x_n, g(x_n))\}$ are the vertices of $\theta(r) \vee g$ and $0 < d_1 < d_2 \dots < d_n$. By linearity then, to show $\int \theta(r) \vee g(x) \mu(dx) \geq \int \theta(r) \vee g(x) \beta(dx)$ it is enough to show $\int y \vee x \mu(dx) \geq \int y \vee x \beta(dx)$ for all $y \in R$.

When $y \in [a, l]$,

$$\begin{aligned} \int x \vee y \mu(dx) &\geq \int x \mu(dx) + (y - a) \mu((-\infty, a]) \\ &\geq c + (y - a)W \\ &= yW + l(1 - W) = \int x \vee y \beta(dy). \end{aligned}$$

When $y \leq a$,

$$\int x \vee y \mu(dx) \geq \int x \mu(dx) \geq c = \int x \vee y \beta(dx).$$

When $y \geq l$,

$$\int x \vee y \mu(dx) \geq y \geq l = \int x \vee y \beta(dx).$$

LEMMA 2.b. With a, c, W, l, μ and β as in Lemma 2.a we have

$$\int_{a^+}^{\infty} \theta(x) \mu(dx) \geq \int_{a^+}^{\infty} \theta(x) \beta(dx)$$

for all convex, increasing functions θ .

Proof. Let $\mu((-\infty, a]) = \tilde{W}$. Define \tilde{l} by $\tilde{W}a + (1 - \tilde{W})\tilde{l} = c$ and let $\tilde{\beta} = \tilde{W}\delta_a + (1 - \tilde{W})\delta_{\tilde{l}} = c$. By Lemma 2.a, taking $r = a$

$$\theta(a) \mu(-\infty, a] + \int_{a^+}^{\infty} \theta(x) \mu(dx) \geq \theta(a) \tilde{\beta}(-\infty, a] + \int_{a^+}^{\infty} \theta(x) \tilde{\beta}(dx)$$

for all convex, increasing functions θ . However $\mu(-\infty, a] = \tilde{\beta}(-\infty, a]$, hence

$$\begin{aligned} \int_{a^+}^{\infty} \theta(x) \mu(dx) &\geq \int_{a^+}^{\infty} \theta(x) \tilde{\beta}(dx) = \theta(\tilde{l})(1 - \tilde{W}) \\ &= (1 - \tilde{W})\theta\left(\frac{c - \tilde{W}(a)}{1 - \tilde{W}}\right) = (1 - \tilde{W})\theta\left(\frac{c - a}{1 - \tilde{W}} + a\right). \end{aligned}$$

However, $(1 - s)\theta((c - a)/(1 - s) + a)$ is an increasing function in $s \in [0, 1]$ (by a supporting hyperplane argument), and $\tilde{W} \geq W$ by Lemma 2.a. Therefore,

$$\begin{aligned} \int_{a^+}^{\infty} \theta(x) \mu(dx) &\geq (1 - \tilde{W})\theta\left(\frac{c - a}{1 - \tilde{W}} + a\right) \\ &\geq (1 - W)\theta\left(\frac{c - a}{1 - W} + a\right) \\ &= \int_{a^+}^{\infty} \theta(x) \beta(dx). \end{aligned}$$

THEOREM 2. *Let P be a submartingale satisfying a condition of type (ϕ, \bar{r}) and such that*

$$\int X_1 dP \geq m.$$

Let the submartingale Q^n on (K^∞, B^∞) be defined by:

$$\begin{aligned} Q^n\{X_1 = m\} &= 1 - q_0 \\ Q^n\{X_1 = a\} &= q_0 - q_1 \\ Q^n\{X_1 = l_1\} &= q_1 \\ Q^n\{X_2 = X_1 | X_1 \neq a\} &= 1 \\ Q^n\{X_2 = b | X_1 = a\} &= 1 \\ (A) \begin{cases} Q^n\{X_3 = X_2 | X_2 \neq b\} &= 1 \\ Q^n\{X_3 = a | X_2 = b\} &= 1 - q_2 \\ Q^n\{X_3 = l_2 | X_2 = b\} &= q_2 \end{cases} \\ (B) \begin{cases} Q^n\{X_4 = X_3 | X_3 \neq a\} &= 1 \\ Q^n\{X_4 = b | X_3 = a\} &= 1 \end{cases} \end{aligned}$$

and so on repeating (A) and (B) for $X_5, X_6, X_7, X_8, \dots, X_{2n-1}, X_{2n}$. $X_{k+1} = X_k$ for $k \geq 2n$. There exist $0 \leq q_0, q_1, \dots, q_n \leq 1$ and $b \leq l_1, \dots, l_n$ such that

- (a) $q_1 = 0$ if $m \leq a$
 $q_1 l_1 + (1 - q_1)a = m, q_0 = 1$ if $m > a$
 $q_k l_k + (1 - q_k)a = b$ if $1 < k \leq n$;
- (b) Q satisfies (ϕ, \bar{r}) ;
- (c) $\int X_1 dQ \geq m$; and
- (d) $P\{T_n\} = Q\{T_n\}$.

(We remark that the trajectories are a.s. $-Q$ of the form (if $m \geq a$):

$$(l_1, l_1, l_1, \dots) \text{ w.p. } q_1;$$

$$(a, b, a, b, \dots a, b, l_r, l_k, \dots) \text{ w.p. } (1 - q_1)(1 - q_2) \dots (1 - q_{k-1})q_k$$

for $1 \leq k \leq n$; and

$$(a, b, a, b, \dots a, b, a, b, \dots) \text{ w.p. } (1 - q_1) \dots (1 - q_n).$$

If $m < a$, the trajectory (l_1, l_1, \dots) is replaced by the trajectory (m, m, \dots) having probability $1 - q_0$.

Proof. We proceed by induction. Suppose the theorem is true for $k \leq n - 1$. Then there exists a submartingale Q^{n-1} of the above form (along with q_0, q_1, \dots, q_{n-1} and l_1, l_2, \dots, l_{n-1}) such that $\int \theta(X_k) dQ^{n-1} \leq \int \theta(X_k) dP$ for $1 \leq k \leq 2n - 2$, and $Q^{n-1}\{T_{n-1}\} = P\{T_{n-1}\}$. Now define $1 - q_n = P\{T_n|T_{n-1}\}$ and l_n by $q_n l_n + (1 - q_n)a = b$. Define Q^n using q_0, q_1, \dots, q_n and l_1, l_2, \dots, l_n . Now consider the probability $\mu(dx) = P\{X_{2n-1} \in dx|T_{n-1}\}$. By Lemma 2.a taking $r = -\infty$, the two point probability $\beta = q_n \delta_{l_n} + (1 - q_n)\delta_a$ satisfies

$$\frac{1}{P\{T_{n-1}\}} \int_{T_{n-1}} \theta(X_{2n-1}) dP = \int \theta(x) \mu(dx) \geq \int \theta(x) \beta(dx)$$

$$= \frac{1}{P\{T_n\}} \int_{T_{n-1}} \theta(X_{2n-1}) dQ^n.$$

Hence

$$(1) \int_{T_{n-1}} \theta(X_{2n-1}) dP \geq \int_{T_{n-1}} \theta(X_{2n-1}) dQ^n.$$

Next,

$$(2) \int_{T_{n-1}^c} \theta(X_{2n-1}) dP \geq \int_{T_{n-1}^c} \theta(X_{2n-2}) dP \geq \int_{T_{n-1}^c} \theta(X_{2n-2}) dQ^{n-1}$$

$$= \int_{T_{n-1}^c} \theta(X_{2n-1}) dQ^n.$$

So

$$\int \theta(X_{2n-1}) dP = \int_{T_{n-1}} \theta(X_{2n-1}) dP + \int_{T_{n-1}^c} \theta(X_{2n-1}) dP$$

$$\geq \int \theta(X_{2n-1}) dQ^n$$

by (1) and (2).

Next,

$$\int \theta(X_{2n})dP = \int_{T_n} \theta(X_{2n})dP + \int_{S_n-T_n} \theta(X_{2n})dP + \int_{T_{n-1}-S_n} \theta(X_{2n})dP$$

$$(3) + \int_{T_{n-1}^c} \theta(X_{2n})dP \geq bP\{T_n\} + \int_{T_{n-1}-S_n} \theta(X_{2n-1})dP + \int_{T_{n-1}^c} \theta(X_{2n-2})dP.$$

Again defining $\mu(dx) = P\{X_{2n-1} \in dx|T_{n-1}\}$,

$$\int_{T_{n-1}-S_n} \theta(X_{2n-1})dP = P\{T_{n-1}\} \int_{a^+}^{\infty} \theta(x) \mu(dx)$$

$$\geq P\{T_{n-1}\} \left(\int_{a^+}^{\infty} \theta(x) \beta(dx) \right) = \int_{T_{n-1}-S_n} \theta(X_{2n-1})dQ^n \text{ (by Lemma 2.b).}$$

Hence from (3),

$$\int \theta(X_{2n})dP \geq bP\{T_n\} + \int_{T_{n-1}-S_n} \theta(X_{2n-1})dQ^n + \int_{T_{n-1}^c} \theta(X_{2n-2})dQ^n$$

$$= \int \theta(X_{2n})dQ^n.$$

Therefore, $\int \theta(X_k)dQ^n \leq \int \theta(X_k)dP$ for $1 \leq k \leq 2n$; $Q^n\{T_n\} = P\{T_n\}$, and by construction, Q^n is a submartingale.

3. Applications. Theorems 1 and 2 provide an algorithm for obtaining sharp upcrossing probabilities. Denote $(x^+)^p$ by $[x]_+^p$.

PROPOSITION 1. *If M is the collection of submartingales such that*

$$\int [X_n - a]_+^p dP \leq L \text{ for all } n, \text{ and}$$

$$\int X_1 dP \geq m, \text{ where } p \geq 1, \text{ and if}$$

$$B_{ab}(k) = \frac{L - [m - a]_+^p}{(b - a)^p} \cdot \frac{(p - 1)^{p-1}}{\left\{ k + \frac{(m - a)^+}{b - a} (p - 1) - 1 \right\}^p + (p - 1)^{p-1}}$$

(if $p = 1$, set $(p - 1)^{p-1} = 1$), then

$$\sup_{P \in M} P\{\gamma_{ab} \geq k\} \leq B_{ab}(k) \text{ for all } k, \text{ and}$$

$$\sup_{P \in M} P\{\gamma_{ab} \geq k\} \sim B_{ab}(k).$$

Proof. Let $\theta_1(x) = [x - a]_+^p$, $\bar{r}(\theta_1) = L$ and $r(\theta_1) = -\infty$. Let $\theta_2(x) = x - m$, $r(\theta_2) = 0$ and $\bar{r}(\theta_2) = \infty$. Let $\phi = \{\theta_1, \theta_2\}$ and \bar{M} be the collection of all submartingales satisfying condition (ϕ, r, \bar{r}) . We check that $\bar{M} = M$. Setting

$B_1 = (-\infty, a], B_2 = [b, \infty), \dots, B_{2n-1} = (-\infty, a], B_{2n} = [b, \infty)$ and applying Theorem 1, we have

$$\sup_{Q \in \tilde{M}} Q\{\gamma_{ab} \geq n\} = \sup_{Q \in \tilde{M}} Q\{X_1 \leq a, \dots, X_{2n} \geq b\}.$$

Next, by Theorem 2,

$$\sup_{Q \in \tilde{M}} Q\{X_1 \leq a, \dots, X_{2n} \geq b\} = \sup_{Q \in \tilde{M}} Q\{X_1 \leq a, \dots, X_{2n} \geq b\},$$

where \tilde{M} is the collection of submartingales in M also having the form given in Theorem 2. Let $Q \in \tilde{M}$. Let

$$Q\{X_1 = m\} = 1 - p_0$$

$$Q\{S_k\} = p_k, \quad k = 1, \dots, n.$$

Therefore, by the submartingale property, $m \leq (1 - p_0)m + p_1a + (p_0 - p_1)l_1$. Clearly equality is best (for satisfying (ϕ, \bar{r})) so

$$l_1 = \frac{p_0m - p_1a}{p_0 - p_1}.$$

Similarly $p_k b = p_{k+1}a + (p_k - p_{k+1})l_{k+1}, k = 1, \dots, n - 1$, so

$$l_{k+1} = \frac{p_k b - p_{k+1}a}{p_k - p_{k+1}}.$$

Next

$$L \geq \int [X_{2n} - a]_+^p dQ$$

$$\geq (1 - p_0)[m - a]_+^p + (p_0 - p_1) \left[\frac{p_0m - p_1a}{p_0 - p_1} - a \right]_+^p$$

$$+ (p_1 - p_2) \left[\frac{p_1b - p_2a}{p_1 - p_2} - a \right]_+^p + \dots$$

$$+ (p_{n-1} - p_n) \left[\frac{p_{n-1}b - p_na}{p_{n-1} - p_n} - a \right]_+^p + p_n(b - a)^p$$

$$= (1 - p_0)[m - a]_+^p + \frac{p_0^p [m - a]_+^p}{(p_0 - p_1)^{p-1}} + \frac{p_1^p}{(p_1 - p_2)^{p-1}} (b - a)^p + \dots$$

$$+ \frac{p_{n-1}^p}{(p_{n-1} - p_n)^{p-1}} (b - a)^p + p_n(b - a)^p.$$

Let

$$\tilde{L} = \frac{L - [m - a]_+^p}{(b - a)^p}, \quad \alpha_0 = p_0, \alpha_1 = \frac{p_1}{p_0}, \dots, \alpha_n = \frac{p_n}{p_{n-1}}.$$

Therefore

$$(1) \quad \tilde{L} \geq \frac{-\alpha_0 [m - a]_+^p}{(b - a)^p} + \frac{\alpha_0 [m - a]_+^p}{(b - a)^p \cdot (1 - \alpha_1)^{p-1}} + \frac{\alpha_0 \alpha_1}{(1 - \alpha_2)^{p-1}} + \dots$$

$$+ \frac{\alpha_0 \dots \alpha_{n-1}}{(1 - \alpha_n)^{p-1}} + \alpha_0 \dots \alpha_n.$$

We must now maximize $p_n = \alpha_0 \dots \alpha_n$ subject to the constraint (1) and $0 \leq \alpha_0, \dots, \alpha_n \leq 1$. Clearly the maximum occurs when (1) is an equality. Solving for α_0 we must maximize

$$(2) \quad \left\{ \frac{\tilde{L} \alpha_1 \dots \alpha_n}{\frac{-(m-a)_+^p}{(b-a)^p} + \frac{[m-a]_+^p}{(b-a)^p(1-\alpha_1)^{p-1}} + \dots + \frac{\alpha_1 \dots \alpha_{n-1}}{(1-\alpha_n)^{p-1}} + \alpha_1 \dots \alpha_n} \right\}.$$

Equivalently, we can minimize

$$\frac{1}{\tilde{L}} \left\{ \frac{[m-a]_+^p}{(b-a)^p \alpha_1 \dots \alpha_n} \left(\frac{1}{(1-\alpha_1)^{p-1}} - 1 \right) + \frac{1}{(1-\alpha_2)^{p-1} \alpha_2 \dots \alpha_k} + \dots + \frac{1}{(1-\alpha_k)^{p-1} \alpha_k} + 1 \right\}.$$

Set

$$\gamma_1 = \frac{[m-a]_+^p}{(b-a)^p \alpha_1} \cdot \left(\frac{1}{(1-\alpha_1)^{p-1}} - 1 \right).$$

For $m \geq 2$ set

$$\gamma_m = \frac{[m-a]_+^p}{(b-a)^p \alpha_1 \dots \alpha_m} \left(\frac{1}{(1-\alpha_1)^{p-1}} - 1 \right) + \frac{1}{(1-\alpha_2)^{p-1} \alpha_2 \dots \alpha_m} + \dots + \frac{1}{(1-\alpha_m)^{p-1} \alpha_m}.$$

Therefore for $k \geq 2$,

$$(3) \quad \gamma_k = \frac{\gamma_{k-1}}{\alpha_k} + \frac{1}{(1-\alpha_k)^{p-1} \alpha_k}.$$

We wish to minimize γ_n by choosing $\bar{\alpha}_1, \dots, \bar{\alpha}_n$. However, γ_{n-1} depends only on $\alpha_1, \dots, \alpha_{n-1}$; therefore, at the minimum,

$$0 = \left. \frac{d\gamma_n}{d\alpha_n} \right|_{\bar{\alpha}_n} = \frac{-\gamma_{n-1}}{\bar{\alpha}_n^2} - \frac{1}{(1-\bar{\alpha}_n)^{p-1} \bar{\alpha}_n^2} + \frac{p-1}{(1-\bar{\alpha}_n)^p \bar{\alpha}_n}.$$

This gives

$$\gamma_{n-1} = \frac{p\bar{\alpha}_n - 1}{(1-\bar{\alpha}_n)^p}.$$

Substituting back into (3) also gives $\gamma_n = (p-1)/(1-\bar{\alpha}_n)^p$. Now at the minimum,

$$0 = \frac{d\gamma_{n-1}}{d\alpha_{n-1}} = \frac{d\gamma_{n-1}}{d\alpha_{n-2}} = \dots = \frac{d\gamma_{n-1}}{d\alpha_1},$$

so the above relations hold for each level. Hence, (henceforth γ_k represents the

minimum value)

$$(4) \quad \gamma_{k-1} = \frac{p\bar{\alpha}_k - 1}{(1 - \bar{\alpha}_k)^p}, \text{ and}$$

$$(5) \quad \gamma_k = \frac{p - 1}{(1 - \bar{\alpha}_k)^p} \text{ for } k \geq 2.$$

We remark that

$$\frac{\tilde{L}}{\gamma_k} \Big|_{\bar{\alpha}_1, \dots, \bar{\alpha}_k} \wedge 1$$

yields the maximum probability of k upcrossings under our constraints for all $k \geq 1$. With our recurrence relations we now examine the asymptotic behavior of γ_k .

$$\frac{\gamma_{k-1}}{\gamma_k} = \frac{p\bar{\alpha}_k - 1}{p - 1} = 1 - \frac{p(1 - \bar{\alpha}_k)}{p - 1}.$$

Also from (5), $(1 - \bar{\alpha}_k) = (p - 1)^{1/p} \gamma_k^{-1/p}$; hence

$$(6) \quad \frac{\gamma_{k-1}}{\gamma_k} = 1 - \frac{p}{(p - 1)^{(p-1)/p}} \gamma_k^{-1/p}, \text{ or}$$

$$(7) \quad \gamma_k - \gamma_{k-1} = C\gamma_k^{1-1/p}; \quad C = \frac{p}{(p - 1)^{(p-1)/p}}.$$

Now consider the equation $d\gamma(t)/dt = C\gamma(t)^{1-1/p}$. Solutions are of the form $\gamma(t) = ((Ct)/p) + C_1)^p$ where C_1 is a constant. Also

$$\begin{aligned} \gamma(k) - \gamma(k - 1) &= \dot{\gamma}(s) \text{ for some } k - 1 \leq s \leq k, \\ &= C\gamma(s)^{1-1/p} \\ &\leq C\gamma(k)^{1-1/p}, \end{aligned}$$

since solutions are increasing. Therefore $\gamma(k)$ increases slower than γ_k .

Next from, (6) we have $\lim_{k \rightarrow \infty} \gamma_{k-1}/\gamma_k = 1$, so for all $\delta > 1$ there exists an n_0 such that for $k \geq n_0$,

$$\gamma_k^{1-1/p} \leq \delta\gamma_{k-1}^{1-1/p} \quad (\delta - 1 \text{ is small}).$$

Hence, $\gamma_k - \gamma_{k-1} \leq \delta C\gamma_{k-1}^{1-1/p}$. Now let $d\tilde{\gamma}(t)/dt = \delta C\tilde{\gamma}(t)^{1-1/p}$. Hence

$$\begin{aligned} \tilde{\gamma}(k) - \tilde{\gamma}(k - 1) &= \dot{\tilde{\gamma}}(s) \text{ for some } k - 1 \leq s \leq k \\ &= \delta C\tilde{\gamma}(s)^{1-1/p} \\ &\geq \delta C\tilde{\gamma}(k - 1)^{1-1/p}. \end{aligned}$$

Therefore $\tilde{\gamma}(k)$ increases faster than γ_k . Also as before $\tilde{\gamma}(t) = ((\delta Ct)/p) + C_2)^p$, where C_2 is a positive constant. Hence γ_k is $o(k^p)$. Moreover, we can solve explicitly for γ_1 .

$$\gamma_1 = \frac{[m - a]_+^p}{(b - a)^p} \frac{1}{\alpha_1} \left(\frac{1}{(1 - \alpha_1)^{p-1}} - 1 \right)$$

is increasing in $0 \leq \alpha_1 \leq 1$. Hence the minimum is

$$\lim_{\alpha_1 \rightarrow 0} \gamma_1 = \frac{[m - a]_+^p}{(b - a)^p} (p - 1).$$

If we set $\gamma(1) = \gamma_1$ we have

$$\left(\frac{C}{p} + C_1\right)^p = \frac{[m - a]_+^p}{(b - a)^p} (p - 1)$$

and after substitution

$$\gamma(t) = \frac{1}{(p - 1)^{p-1}} \left(t + \frac{(m - a)^+}{(b - a)} (p - 1) - 1\right)^p.$$

Hence

$$\begin{aligned} \frac{\tilde{L}}{\gamma_k + 1} &\leq \frac{\tilde{L}}{\gamma(k) + 1} = \frac{L - [m - a]_+^p}{(b - a)^p} \\ &\quad \times \frac{(p - 1)^{p-1}}{\left(k + \frac{(m - a)^+}{b - a} (p - 1) - 1\right)^p + (p - 1)^{p-1}}. \end{aligned}$$

Since the probability of k upcrossings is at most $\tilde{L}/(\gamma_k + 1)$ we have our bound.

We now set $\tilde{\gamma}(n_0) = \gamma_{n_0}$, thereby determining C_2 . Hence

$$\begin{aligned} \frac{\gamma_k}{\gamma(k)} &\leq \frac{\tilde{\gamma}(k)}{\gamma(k)} \quad \text{for } k \geq n_0 \\ &\leq \frac{\left(\frac{\delta Ck}{p} + C_2\right)^p}{\left(\frac{Ck}{p} + C_1\right)^p}. \end{aligned}$$

Therefore

$$\overline{\lim}_{k \rightarrow \infty} \frac{\gamma_k}{\gamma(k)} \leq \delta^p;$$

but $\delta - 1$ is arbitrarily small. Therefore

$$\overline{\lim}_{k \rightarrow \infty} \frac{\gamma_k}{\gamma(k)} \leq 1.$$

Hence

$$\lim_{k \rightarrow \infty} \frac{\frac{\tilde{L}}{\gamma(k) + 1}}{\frac{\tilde{L}}{\gamma_k + 1}} = 1,$$

so

$$\sup_{P \in \mathcal{M}} P\{\gamma_{ab} \geq k\} \sim B_{ab}(k).$$

For $p = 1$ it is easiest to maximize the expression (2). The maximum is \bar{L}/k . This completes the proof.

We could generalize Proposition 1 by supposing $EX_1 = m$ and $E\theta(X_n - a) \leq L$ for all n where θ is an increasing convex function with derivative θ' . The above proof goes through and (7) becomes $\gamma_k - \gamma_{k-1} = \theta' \circ \alpha(\gamma_k)$ where $\alpha \circ (\theta'(x)x - \theta(x)) = x$. In a particular case we may be able to proceed (as above) from here.

COROLLARY 1. *Doob's upcrossing inequality is sharp.*

Proof. For any submartingale P such that $\int (X_n - a)^+ dP \leq L$ for all n and $\int X_1 = m$, Doob's inequality says

$$\int \gamma_{ab} dP \leq \frac{L - (m - a)^+}{b - a}.$$

Applying Chebyshev's inequality, we have

$$P\{\gamma_{ab} \geq k\} \leq \frac{L - (m - a)^+}{k(b - a)}$$

which is precisely the bound given in Proposition 1 for $p = 1$ (Prof. David Heath pointed this out). Proposition 1 provides the construction of a submartingale (almost) attaining this bound (in fact $\alpha_2 = \alpha_3 = \dots = \alpha_k = 1$ means $l_1 = l_2 = \dots = l_k = \infty$ so at best by taking l_1, \dots, l_k large we may almost attain the bound). Hence Doob's inequality must also be sharp.

It is in fact possible to obtain Doob's upcrossing inequality directly by these methods (see [2]).

Example 2. (Dubins' inequality—see [3, p. 27]).

PROPOSITION 2. *If P is a submartingale such that $P\{L \leq X_n \leq U\} = 1$ for constants L, U ($L \leq U$) for all n , and $\int X_1 dP \geq m$, then*

$$P\{\gamma_{ab} \geq k\} \leq \left(\frac{U - m \vee a}{U - a} \right) \left(\frac{U - b}{U - a} \right)^{k-1}$$

for $L \leq a < b \leq U$, where $m \vee a = \max\{m, a\}$.

Proof. Define $\theta(x) = x$, $r(\theta) = m$, $\bar{r}(\theta) = \infty$ and $\phi = \{\theta\}$. Let M be the class of all submartingales satisfying conditions (ϕ, r, \bar{r}) and (L, U) . It is clear that M consists of exactly those submartingales satisfying our hypotheses. Therefore by Theorem 1,

$$\sup_{P \in \mathcal{M}} P\{\gamma_{ab} \geq k\} = \sup_{P \in \mathcal{M}} P\{T_k\}.$$

For any $P \in M$,

$$\int_{T_{k-1}} X_{2k-2} dP \leq \int_{T_{k-1}} X_{2k-1} dP = \int_{S_k} X_{2k-1} dP + \int_{T_{k-1} \cap S_k^c} X_{2k-1} dP.$$

So $bP\{T_{k-1}\} \leq P\{S_k\}a + (P\{T_{k-1}\} - P\{S_k\})U$. Hence

$$P\{T_k\} \leq P\{S_k\} \leq \frac{U-b}{U-a} P\{T_{k-1}\}.$$

Next $m \leq \int X_1 dP = \int_{S_1} X_1 dP + \int_{S_1^c} X_1 dP$, so $m \leq P\{S_1\}a + (1 - P\{S_1\})U$. Hence

$$P\{T_1\} \leq P\{S_1\} \leq \left(\frac{U-m}{U-a}\right) \wedge 1.$$

By iteration we have

$$P\{T_k\} \leq \left(\frac{U-m \vee a}{U-a}\right) \left(\frac{U-b}{U-a}\right)^{k-1}.$$

Again this bound is sharp. A martingale with precisely these upcrossing probabilities is given as an Exercise II-2 in [3].

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