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Behaviour of the minimum degree throughout the *d***-process**[∗]

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Abstract

The *d*-process generates a graph at random by starting with an empty graph with *n* vertices, then adding edges one at a time uniformly at random among all pairs of vertices which have degrees at most *d* − 1 and are not mutually joined. We show that, in the evolution of a random graph with *n* vertices under the *d*-process with *d* fixed, with high probability, for each $j \in \{0, 1, \ldots, d-2\}$, the minimum degree jumps from *j* to *j* + 1 when the number of steps left is on the order of $\ln(n)^{d-j-1}$. This answers a question of Rucinski and Wormald. More specifically, we show that, when the last vertex of degree *j* disappears, the number of steps left divided by ln (*n*)^{*d−j*−1} converges in distribution to the exponential random variable of mean $\frac{j!}{2(d-1)!}$; furthermore, these *d* − 1 distributions are independent.

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1. Introduction

There are numerous models that generate different types of sparse random graphs. Among them is the *d*-*process*, defined in the following way: start with *n* vertices and 0 edges, and at each time step, choose a pair $\{u, v\}$ uniformly at random over all pairs consisting of vertices which have degree less than *d* and are not joined to each other by an edge. *d* could be allowed to change with *n*, but for the rest of this paper *d* is always a fixed constant (this is also assumed in all relevant citations). Rucinski and Wormald showed that with high probability, abbreviated "w.h.p." (i.e. ´ with probability converging to 1 as $n \to \infty$) the *d*-process ends with $|dn/2|$ edges [\[11\]](#page-18-0). There is still much that is unknown about the *d*-process; for example, it is not known whether the *d*process is *contiguous* with the *d*-uniform random graph model for any $d \geq 2$; i.e. if any event that happens with high probability in one happens with high probability in the other. A recent paper by Molloy, Surya, and Warnke [\[8\]](#page-18-1) disproves this relation if there is enough "non-uniformity" of the vertex degrees (with an appropriate modification of the *d*-process for non-regular graphs); it also contains a good history of the *d*-process. See ref. [\[7,](#page-18-2) Section 9.6] for more on contiguity.

A couple of notable results have been given for the case where $d = 2$: the expected numbers of cycles of constant sizes were studied by Ruciński and Wormald in ref. [[10\]](#page-18-3), and in ref. [\[13\]](#page-18-4), Telcs, Wormald, and Zhou calculated the probability that the 2-process ends with a Hamiltonian cycle.

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In these works, the authors establish estimates on certain graph parameters, such as the number of vertices of a certain degree, that hold throughout the process. This is done with the so-called "differential equations method" for random graph processes, which uses martingale inequalities to give variable bounds; in ref. [\[14\]](#page-18-5) Wormald gives a thorough description of this method.

More recently, Rucinski and Wormald announced a new analysis of the d-process that hinges on a coupling with a balls-in-bins process. This simple argument gives a precise estimate of the probability that the *d*-process ends with $|dn/2|$ edges (i.e. the probability that the *d*-process reaches saturation). This argument includes estimaes for the number of vertices of each degree near the end of the process. This work was presented by Rucinski at the 2023 *Random Structures and Algorithms* conference. Ruciński's presentation included the following problem (which was open at the time): when do we expect the last vertex of degree *j* (for any *j* from 0 to $d - 2$) to disappear? The question was also stated earlier for $d = 2$ and $j = 0$ by Rucinski and Wormald [[10,](#page-18-3) Question 3]. In November of 2023, after the first release of the pre-print of this paper, Rucinski ´ and Wormald released a pre-print of their balls-and-bins argument which also included an answer to Rucinski's question $[12]$ $[12]$. Our main result uses the differential equations method (as described in the previous paragraph) and gives a slightly stronger answer:

Theorem 1. *Consider the d-process on a vertex set of size n, and for each* $\ell \in \{0\} \cup [d-2]$ *, let the* random variable T_ℓ be the step at which the number of vertices of degree at most ℓ becomes 0. Then *the sequence (over n) of random d* − 1*-tuples consisting of the variables*

$$
V_n^{(\ell)} = \frac{(d-1)!(dn - 2T_{\ell})}{\ell! (\ln(n))^{d-1-\ell}}
$$

converges in distribution to the product of d − 1 *independent exponential random variables of mean 1.*

In this paper we use the differential equations method with increasingly precise estimates of certain random variables; these estimates are known as *self-correcting*. Previous results that use self-correcting estimates include [\[13\]](#page-18-4), [\[6\]](#page-18-7), [\[3\]](#page-17-0), [\[4\]](#page-17-1), [\[5\]](#page-17-2), and [\[9\]](#page-18-8). There have been various approaches to achieving self-correcting estimates; the approach in this paper uses *critical intervals*, regions of possible values of a random variable in which we expect subsequent variables to increase/decrease over time. Critical intervals used in this fashion first appeared in a result of Bohman and Picollelli [\[6\]](#page-18-7). For an introduction to and discussion of the method see Bohman, Frieze, and Lubetzky [\[3\]](#page-17-0).

The proof of Theorem [1](#page-1-0) is divided into four sections. In Section [2,](#page-2-0) we introduce random variables of the form $S_i^{(j)}$ which count the number of vertices of degree at most *j* after *i* steps, define approximating functions $s_j(t)$ with the eventual goal of showing that $S_i^{(j)} \approx n s_j(i/n)$ for most of the process, and derive useful properties of these functions. One such property is that, when there are at most n^c steps left for some constant $c < 1$,

$$
\frac{s_j(i/n)}{s_{j-1}(i/n)} = \Theta(\ln(n))
$$

for each *j*; this hierarchy of functions helps us to focus on each variable $S_i^{(j)}$ independently of the others when it is near 0, which motivates the form of the limiting exponential random variables in Theorem [1.](#page-1-0) At the end of Section [2](#page-2-0) we introduce two martingale inequalities used by Bohman [\[2\]](#page-17-3) and make a slight modification to them to use later in the paper. In Section [3,](#page-5-0) we work with a more 'standard martingale method' (without the use of critical intervals) to show that $S_i^{(j)} \approx ns_j(i/n)$ until there are $n^{1-1/(100d)}$ steps left. Here we allow the error bounds to increase over time. In Section [4,](#page-7-0) we use a more refined martingale method (including the use of critical intervals) to show that $S_i^{(j)}$ ≈ $ns_j(i/n)$ continues to hold until there are ln (*n*)^{*d*−0.8−*j* steps left; here the error bounds}

decrease over time, and so are self-correcting. In Section [5,](#page-13-0) we complete the proof of Theorem [1](#page-1-0) by using a pairing argument to show that, in the last steps of the *d*-process, the behaviour of the random variable in question can be well-approximated by a certain uniform distribution of time steps. Sections [4](#page-7-0) and [5](#page-13-0) are both parts of a proof by induction over a series of intervals of time steps, though we give each part its own section as the methods used in each are very different.

2. Preliminaries

First, two technical notes: we use the standard notation of symbols $o, O, \Theta, \omega, \Omega, \ll, \gg$, and \sim to compare functions asymptotically (e.g. see pages 9-10 of [\[7\]](#page-18-2)). We also note that, throughout the paper, we assume *n* to be arbitrarily large.

In this section we set up sequences of random variables, describe how the evolution of the *d*-process depends on these, and deduce properties of certain *approximating functions*; such functions are used to estimate the number of vertices of given degrees throughout the process (much of this is also described in ref. [\[13\]](#page-18-4) with similar notation; the one major difference is that we use *i* instead of *t* for the number of time steps, and *t* instead of *x* for the corresponding time variable). Consider a sequence of graphs $G_0, G_1, \ldots, G_{\lfloor dn/2 \rfloor}$, where G_0 is the empty graph of *n* vertices, and for *i* ∈ [*n*], let *Gi* be formed by adding an edge uniformly at random to *Gi*[−]¹ so that the maximum degree of *Gi* is at most *d* (in the unlikely event that there are no valid edges to add after *s* steps for some $s < \lfloor d_n/2 \rfloor$, let $G_i = G_s$ for all $i > s$). Next, we define several sequences of random variables: For all *i*, *j*, *j*₁, *j*₂ such that $0 \le i \le \lfloor \frac{dn}{2} \rfloor$, $0 \le j \le d$, and $0 \le j_1 \le j_2 \le d - 1$, define:

 $Y^{(j)}_{i}$:= the number of vertices in G_i with degree *j* $S_i^{(j)} :=$ the number of vertices in G_i with degree *at most j* $Z_i^{(j_1,j_2)} :=$ the number of edges $\{v_1, v_2\}$ in G_i for which

 $\min\{\text{deg}(v_1), \text{deg}(v_2)\} = j_1$ and $\max\{\text{deg}(v_1), \text{deg}(v_2)\} = j_2$

$$
Z_i := \sum_{0 \le j_1 \le j_2 \le d-1} Z_i^{(j_1, j_2)}.
$$

By definition and by edge-counting we have

$$
\sum_{j=0}^{d} Y_i^{(j)} = n \quad \text{and} \quad \sum_{j=1}^{d} j Y_i^{(j)} = 2i.
$$

Combining the equations above gives us

$$
\sum_{j=0}^{d-1} S_i^{(j)} = \sum_{j=0}^{d-1} (d-j) Y_i^{(j)} = dn - 2i.
$$
 (1)

One can also quickly verify from [\(1\)](#page-2-1) and by definition of $S_i^{(d-1)}$, $Z_i^{(j_1,j_2)}$, and Z_i , that

$$
dS_i^{(d-1)} \ge \max\{2Z_i, dn-2i\} \quad \text{and} \quad jY_i^{(j)} \ge \sum_{k \le j} Z_i^{(k,j)} + \sum_{k \ge j} Z_i^{(j,k)}.
$$
 (2)

Throughout the process we will keep track of the variables $S^{(j)}$ using martingale arguments. This is sufficient for us, as the *Y* variables can be derived from the *S* variables, and because none of the *Z* variables will have any significant effect in any of our calculations, as we will see later.

Our next step is to estimate the expected on-step change of $S_i^{(j)}$, known as the "trend hypothesis" in ref. [\[14\]](#page-18-5). Note that $S_i^{(j)} - S_{i+1}^{(j)}$ equals the number of vertices of degree *j* that are picked at the *i* + 1 time step; hence, for all *j* ∈ {0} ∪ [*d* − 1]:

$$
\mathbb{E}\left[S_{i+1}^{(j)} - S_i^{(j)} | G_i\right] = \frac{-Y_i^{(j)} \left(S_i^{(d-1)} - 1\right) + \sum_{k \le j} Z_i^{(k,j)} + \sum_{k \ge j} Z_i^{(j,k)}}{\binom{S_i^{(d-1)}}{2} - Z_i}
$$
\n
$$
= \frac{-2Y_i^{(j)}}{S_i^{(d-1)}} \left(1 + O\left(\frac{1}{dn - 2i}\right)\right) \qquad \text{by (2)}
$$
\n(3)

$$
=\frac{-2Y_i^{(j)}}{S_i^{(d-1)}}+O\left(\frac{1}{dn-2i}\right).
$$
\n(4)

For $j \in \{0\} \cup [d-1]$ we define approximating functions $y_j:[0, d/2) \to \mathbb{R}$ and $s_j:[0, d/2) \to \mathbb{R}$: let *y_j*(*t*), *s_j*(*t*) be functions such that $s_j = \sum_{k=0}^{j} y_k$, $y_0(0) = 1$ and $y_k(0) = 0$ for all $k \in [d-1]$ (equivalent to $s_j(0) = 1$ for all *j*), and (assuming the "dummy functions" $y_{-1}(t) = s_{-1}(t) = 0$):

$$
\frac{ds_j}{dt} = \frac{-2y_j}{s_{d-1}} = \frac{2(s_{j-1} - s_j)}{s_{d-1}} \qquad \frac{dy_j}{dt} = \frac{2(y_{j-1} - y_j)}{s_{d-1}}.
$$
(5)

By [\(5\)](#page-3-0) and the chain rule, for all $j \in [d-1]$:

$$
\frac{dy_j}{dy_0} - \frac{y_j}{y_0} = -\frac{y_{j-1}}{y_0}.
$$

Since the above equation is first-order linear, we have, for some constant *Cj*:

$$
y_j = y_0 \left(- \int \frac{y_{j-1}}{y_0^2} dy_0 + C_j \right).
$$

Using the above recursively with initial conditions, we have, for all $j \in [d-1]$:

$$
y_j = \frac{y_0(-\ln(y_0))^j}{j!}.
$$
 (6)

To solve for an explicit formula relating y_0 and *t*, note that, by [\(5\)](#page-3-0):

$$
\frac{d}{dt}\left(\sum_{j=0}^{d-1}\left(d-j\right)y_j\right) = \frac{-2\sum_{j=0}^{d-1}y_j}{s_{d-1}} = -2,
$$

hence, using initial conditions (note the resemblance to [\(1\)](#page-2-1)):

$$
\sum_{j=0}^{d-1} s_j = \sum_{j=0}^{d-1} (d-j)y_j = d - 2t.
$$
 (7)

Equations [\(6\)](#page-3-1) and [\(7\)](#page-3-2) together give us a complete description of the functions y_i and s_j . We will now prove some useful properties of these functions. To start, we can combine [\(6\)](#page-3-1) and [\(7\)](#page-3-2) to get

$$
\sum_{j=0}^{d-1} \frac{y_0(-\ln(y_0))^j (d-j)}{j!} = d - 2t.
$$
 (8)

Note that, by continuity of y_0 and by [\(8\)](#page-3-3), $y_0 > 0$ over its domain. Next, by summing up [\(6\)](#page-3-1) over *j* ∈ {0} ∪ [*d* − 1] ([\(6\)](#page-3-1) holds for *j* = 0 also) one can see that s_{d-1} is positive if y_0 < 1. This, combined with $\frac{dy_0}{dt} = \frac{-2y_0}{s_{d-1}}$, tells us that y_0 is decreasing and s_{d-1} is positive. It follows from $y_0 \in (0, 1]$ and [\(6\)](#page-3-1) that each y_j is positive for *t* ≠ 0. In turn, this implies that $0 \le y_j \le s_j \le s_{d-1}$ for each *j*. From this it follows that ds_{d-1} is at least the left expression of [\(7\)](#page-3-2), so $s_{d-1} \geq 1 - \frac{2t}{d}$. We make a special note of the last couple of properties mentioned:

$$
0 \le y_j \le s_j \le s_{d-1}
$$
 for all j and $s_{d-1}(i/n) \ge 1 - \frac{2i}{dn}$. (9)

Next, we want to understand the behaviour of each function when *t* is close to $\frac{d}{2}$, as this is the most critical point of the process. Consider [\(8\)](#page-3-3) again. As $t \to \frac{d}{2}$, $y_0 \to 0$, so $\frac{y_0(-\ln(y_0))^{d-1}}{(d-1)!}$ will be the most dominant term on the left; hence,

$$
t \to \frac{d}{2} \implies y_0 \sim \frac{(d-1)!(d-2t)}{(-\ln\left(d-2t\right))^{d-1}}.
$$

This, combined with [\(6\)](#page-3-1) gives us, for all $j \in \{0\} \cup [d-1]$:

$$
t \to \frac{d}{2} \implies y_j(t) \sim s_j(t) \sim \frac{(d-1)!(d-2t)}{j!(-\ln(d-2t))^{d-1-j}}.
$$
\n(10)

For large enough *t* (and hence, for a large enough step *i*), we can approximate the above expression:

$$
i \ge \frac{dn}{2} - n^{1-1/(100d)} \implies ny_j(i/n) \sim ns_j(i/n) = \Theta\left(\ln\left(n\right)^{-d+1+j}\left(\frac{dn}{2} - i\right)\right). \tag{11}
$$

One can now see that, near the end of the process, $s_i/s_{i-1} = \Theta(\ln(n))$, as mentioned in the introduction.

Finally, we introduce two martingale inequalities from a result of Bohman [\[2\]](#page-17-3) which will be used in Section [4](#page-7-0) in a slightly modified form. The original inequalities are as follows:

Lemma 2 (Lemma 6 from [\[2\]](#page-17-3))*. Suppose a,* η *, and N are positive,* $\eta \le N/2$ *, and a* < ηm *. If* 0 = A_0, A_1, \ldots, A_m *is a submartingale such that* $-\eta \leq A_{i+1} - A_i \leq N$ *for all i, then*

$$
\mathbb{P}[A_m \leq -a] \leq e^{-\frac{a^2}{3\eta N m}}.
$$

Lemma 3 (Lemma 7 from [\[2\]](#page-17-3))*. Suppose a, η, and N are positive,* $\eta \leq N/10$ *, and a < ηm. If* 0 = A_0, A_1, \ldots, A_m *is a supermartingale such that* $-\eta \leq A_{i+1} - A_i \leq N$ for all *i*, then

$$
\mathbb{P}[A_m \geq a] \leq e^{-\frac{a^2}{3\eta N m}}.
$$

We present the following modification, which removes the requirement $a < \eta m$ and modifies one of the inequalities slightly:

Corollary 4. *Suppose a,* η *, and N are positive, and* $\eta \le N/2$ *. If* $0 = A_0, A_1, \ldots, A_m$ *is a submartingale such that* $-\eta \leq A_{i+1} - A_i \leq N$ *for all i, then*

$$
\mathbb{P}[A_m \leq -a] \leq e^{-\frac{a^2}{3\eta N m}}.
$$

Corollary 5. Suppose a, η , and N are positive, and $\eta \le N/10$. If $0 = A_0, A_1, \ldots, A_m$ is a super*martingale such that* $-\eta \leq A_{i+1} - A_i \leq N$ *for all i, then*

$$
\mathbb{P}[A_m \geq a] \leq e^{-\frac{a^2}{3\eta N m}} + e^{-\frac{a}{6N}}.
$$

Corollary [4](#page-4-0) is nearly immediate from Lemma [2:](#page-4-1) first, one can extend the result to include $a =$ η*m* by using left-continuity (with respect to *a*) of both sides of the inequality; we hence assume *a* > *nm*. Since $A_{i+1} - A_i \ge -\eta$ > $-a/m$, then $A_m = A_m - A_0$ > $-a$. We now derive Corollary [5](#page-4-2) from Lemma [3:](#page-4-3) assume $a \ge \eta m$, and let $m' \in \mathbb{Z}^+$ such that $a < \eta m' \le 2a$. Extend the martingale by adding variables $A_{m+1}, \ldots, A_{m'}$ which are all equal to A_m . Apply Lemma [3](#page-4-3) with *m* replaced with *m'*, and use $\eta m' \leq 2a$ to get

$$
\mathbb{P}[A_m \geq a] = \mathbb{P}[A_{m'} \geq a] \leq e^{-\frac{a}{6N}}.
$$

Combining the case $a < \eta m$ from Lemma [3](#page-4-3) and the case $a \geq \eta m$ above gives Corollary [5.](#page-4-2)

3. First phase

Let $i_{trans} = \lfloor \frac{dn}{2} - n^{1-1/(100d)} \rfloor$. The objective of this section is to prove the following Theorem:

Theorem 6. Define $E_{first}(i) := n^{0.6} \left(\frac{dn}{dn-2i}\right)^{4d}$. With high probability, for all $i \leq i_{trans}$ and all $j \in$ {0} ∪ [*d* − 1]*:*

$$
\left| S_i^{(j)} - n s_j \left(\frac{i}{n} \right) \right| \le E_{\text{first}}(i). \tag{12}
$$

Now we define two new random variables for each *j*:

$$
S_i^{(j)+} := S_i^{(j)} - ns_j(i/n) - E_{first}(i)
$$

$$
S_i^{(j)-} := S_i^{(j)} - ns_j(i/n) + E_{first}(i).
$$

Next, we introduce a stopping time *T*, defined as the first step $i \leq i_{trans}$ for which [\(12\)](#page-5-1) is *not* satisfied for some *j*; if [\(12\)](#page-5-1) always holds, then let $T = \infty$. Although this stopping time is not necessarily needed to prove Theorem [6,](#page-5-2) it does make some calculations easier, and moreover, a similar stopping time *will* be necessary for the following section; hence, this serves as a good warm-up. Let variable name *W* be introduced to equip this stopping time to variable *S*, i.e.

$$
W_i^{(j)+} := \begin{cases} S_i^{(j)+}, i < T \\ S_T^{(j)+}, i \ge T \end{cases} \qquad W_i^{(j)-} := \begin{cases} S_i^{(j)-}, i < T \\ S_T^{(j)-}, i \ge T. \end{cases}
$$

Note that $W_i^{(j)+}$ corresponds to the upper boundary and $W_i^{(j)-}$ to the lower one in the sense that crossing the corresponding boundary will make the corresponding variable change signs; $f(x)$ furthermore, the inequality of Theorem [6](#page-5-2) holds if and only if $W_{i_{trans}}^{(j)+}$ ≤ 0 and $W_{i_{trans}}^{(j)-}$ ≥ 0 for each *j*. We now state our martingale Lemma:

Lemma 7. Restricted to $i \le i_{trans}$, for all j, $(W_i^{(j)-})_i$ is a submartingale and $(W_i^{(j)+})_i$ is a super*martingale.*

Proof. Here we just prove the first part of the Lemma; the second part follows from nearly identical calculations. Fix some arbitrary $i \leq i_{trans}$; we need to show that

$$
\mathbb{E}\left[W_{i+1}^{(j)-}-W_i^{(j)-}\mid G_i\right]\geq 0.
$$

Also assume that $T \ge i + 1$, else $W_{i+1}^{(j)-} - W_i^{(j)-} = 0$ and we are done; it follows that $W_i^{(j)-} = 0$ $S_i^{(j)-}$, $W_{i+1}^{(j)-} = S_{i+1}^{(j)-}$, and [\(12\)](#page-5-1) holds for the fixed *i*. By [\(4\)](#page-3-4) and [\(5\)](#page-3-0) and using Taylor's Theorem, we have, for some $\psi \in [i, i + 1]$:

$$
\mathbb{E}\left[S_{i+1}^{(j)}-S_i^{(j)}\mid G_i\right] = \frac{-2Y_i^{(j)}}{S_i^{(d-1)}} + O\left(\frac{1}{dn-2i}\right) + \frac{2y_j(i/n)}{s_{d-1}(i/n)} - \frac{d^2}{d\mu^2} \left(\frac{ns_j(\mu/n)}{2}\right)\Big|_{\mu=\psi} + \left(E_{first}(i+1) - E_{first}(i)\right).
$$

We split the above expression $\left(\text{excluding } O\left(\frac{1}{dn-2i}\right)\right)$ into three summands.

1. Here we make use of the fact that $Y_i^{(j)} = S_i^{(j)} - S_i^{(j-1)}$ and $y_j(t) = s_j(t) - s_{j-1}(t)$. We have (putting s_{j-1} and $S_i^{(-1)} = 0$):

$$
\frac{-2Y_i^{(j)}}{S_i^{(d-1)}} + \frac{2y_j(i/n)}{s_{d-1}(i/n)} = \frac{-2S_i^{(j)} + 2ns_j(i/n) + 2S_i^{(j-1)} - 2ns_{j-1}(i/n)}{ns_{d-1}(i/n)}
$$

+
$$
2Y_i^{(j)} \left(\frac{1}{ns_{d-1}(i/n)} - \frac{1}{S_i^{(d-1)}}\right)
$$

$$
\geq \frac{-4E_{first}(i)}{ns_{d-1}(i/n)} + 2Y_i^{(j)} \left(\frac{1}{ns_{d-1}(i/n)} - \frac{1}{S_i^{(d-1)}}\right) \qquad \text{by (12) and } i < T
$$

$$
\geq \frac{-4E_{first}(i)}{ns_{d-1}(i/n)} - \frac{2Y_i^{(j)}E_{first}(i)}{S_i^{(d-1)}(ns_{d-1}(i/n))} \qquad \text{by (12) and } i < T
$$

$$
\geq \frac{-6E_{first}(i)}{ns_{d-1}(i/n)}.
$$

2.

$$
-\frac{d^2}{d\mu^2} \left(\frac{ns_j(\mu/n)}{2}\right) \Big|_{\mu=\psi} = \frac{2}{n} \left(\frac{s_{d-1}(\psi/n)(y_{j-1}(\psi/n) - y_j(\psi/n)) + y_j(\psi/n)y_{d-1}(\psi/n)}{(s_{d-1}(\psi/n))^3}\right)
$$

$$
= O\left(\frac{1}{dn-2i}\right) \qquad \text{by (9)}.
$$

3. For some $\phi \in [i, i + 1]$:

$$
E_{\text{first}}(i+1) - E_{\text{first}}(i) = \frac{dE_{\text{first}}(\mu)}{d\mu}\Big|_{\mu = \phi}
$$

= $8d^{4d+1}n^{4d+0.6} (dn - 2\phi)^{-4d-1}$
= $(1 + o(1)) \frac{8dE_{\text{first}}(i)}{dn - 2i}.$ (13)

 $\hfill \square$

Now we put the three bounds together:

$$
\mathbb{E}\left[Y_{i+1}^{-} - Y_i^{-} \mid G_i\right] \ge \frac{7dE_{first}(i)}{dn - 2i} - \frac{6E_{first}(i)}{ns_{d-1}(i/n)} + O\left(\frac{1}{dn - 2i}\right)
$$

$$
\ge \frac{dE_{first}(i) + O(1)}{dn - 2i} \qquad \text{by (9)}
$$

$$
\ge 0.
$$

Next, we need a Lipschitz condition on each of our variables. Note that $S_{i+1}^{(j)} - S_i^{(j)}$ is either -2 , −1, or 0; also, one can quickly verify that $|s_j((i+1)/n) - s_j(i/n)| \leq \frac{2}{n}$ by [\(5\)](#page-3-0) and [\(9\)](#page-4-4), and $\left| E_{\text{first}}(i+1) - E_{\text{first}}(i) \right| = o(1)$ by [\(13\)](#page-6-0). Hence, we have, for all $i \leq i_{\text{trans}}$ and all *j*:

$$
\max\left\{ \left| W_{i+1}^{(j)+} - W_i^{(j)+} \right|, \left| W_{i+1}^{(j)-} - W_i^{(j)-} \right| \right\} \le 5. \tag{14}
$$

We conclude the proof of Theorem [6](#page-5-2) by noting that, by Lemma [7](#page-5-3) and (14) , we can use the standard Hoeffding-Azuma inequality for martingales (e.g. Theorem 7.2.1 in [\[1\]](#page-17-4)) to show that $\mathbb{P}\left[\, W^{(j)+}_{i_{trans}} > 0 \, \right]$ and $\mathbb{P}\left[\, W^{(j)-}_{i_{trans}} < 0 \, \right]$ are both $o(1).$ For example, for the variable $W^{(j)+}_{i}$ one would get

$$
\mathbb{P}\left[W_{i_{trans}}^{(j)+}>0\right]\leq \exp\left\{-\frac{n^{1.2}}{50i_{trans}}\right\}=o(1).
$$

4. Second phase

The second phase is where the more sophisticated tools will be used, including the use of critical intervals, self-correcting estimates, and a more general martingale inequality. Furthermore, this phase is broken up into *d* − 1 sub-phases, in relation to when each of the *d* − 1 sequences *S*(*j*) (for *j* ≤ *d* − 2) terminate at 0. First, a few definitions: for all *k* ∈ {0} ∪ [*d* − 2], define

$$
i_{after}(k) := \left\lfloor \frac{dn}{2} - \ln{(n)}^{d-1.01-k} \right\rfloor.
$$

These step values will govern the endpoints of the sub-phases: define for all $k \in \{0\} \cup [d-2]$:

$$
I_k := \begin{cases} \left[i_{trans} + 1, i_{after}(0)\right], & k = 0\\ \left[i_{after}(k-1) + 1, i_{after}(k)\right], & k > 0. \end{cases}
$$

Next, for all *i*, *j*, *k* such that $0 \leq j < d$, $0 \leq k < d - 1$, and $i \in I_k$, define error functions

$$
E_{j,k}(i) = E_j(i) := 2^k \ln(n)^{0.05} (n s_j(i/n))^{0.7}.
$$

Note that, by [\(11\)](#page-4-5), we have

$$
E_j(i) = \Theta\left(\ln\left(n\right)^{-0.7d+0.75+0.7j}\left(\frac{dn}{2}-i\right)^{0.7}\right).
$$
\n(15)

Finally, for any $r \in \mathbb{R}_+$ and $\ell \in [d-2]$, define

$$
i(r,\ell) = \frac{dn}{2} - \left(\frac{\ell!}{2(d-1)!}\right) r(\ln n)^{d-1-\ell}.
$$

The following Theorem will be proved by induction over the *d* − 1 sub-phases governed by the index *k*:

Theorem 8. *For each* $k \in \{0\}$ ∪ $[d - 2]$ *:*

1. With high probability, for all integers $j \in [0, d − 1]$ and $i \in I_k$:

$$
\left| S_i^{(j)} - n s_j \left(\frac{i}{n} \right) \right| \le 4 E_j(i). \tag{16}
$$

2. $S^{(k)}_{i_{after}(k)} = 0$ *with high probability. Furthermore, for any k* + 1*-tuple* {*r*₀, *r*₁, ..., *r*_{*k*}} ∈ (R⁺ ∪ {0}) *^k*+1*:*

$$
\mathbb{P}\left(\bigcap_{\ell=0}^k \left(S^{(\ell)}_{\lfloor i(r_\ell,\ell)\rfloor}=0\right)\right) \to \exp\left\{-\sum_{\ell=0}^k r_\ell\right\}.
$$

In the end, it is only the second statement with $k = d - 2$ that matters for proving Theorem [1.](#page-1-0) We make the connection here:

Proof of Theorem [1](#page-1-0) from Theorem [8.](#page-7-2) First, note that $S_{\lfloor i(r_\ell,\ell) \rfloor}^{(\ell)} = 0$ is the same as $T_\ell \leq i(r_\ell,\ell)$, hence by Theorem [8:](#page-7-2)

$$
\mathbb{P}\left(\bigcap_{\ell=0}^{d-2} \left(T_{\ell} \leq i(r_{\ell}, \ell)\right)\right) \to \exp\left\{-\sum_{\ell=0}^{d-2} r_{\ell}\right\}.
$$

Using the Principle of Inclusion-Exclusion plus a simple limiting argument, one can derive

$$
\mathbb{P}\left(\bigcap_{\ell=0}^{d-2}\left(\frac{(d-1)!(dn-2T_{\ell})}{\ell!(\ln(n))^{d-1-\ell}}\leq r_{\ell}\right)\right)=\mathbb{P}\left(\bigcap_{\ell=0}^{d-2}\left(T_{\ell}\geq i(r_{\ell},\ell)\right)\right)\to\prod_{\ell=0}^{d-2}\left(1-e^{-r_{\ell}}\right),
$$

hence the *d* − 1-dimensional random vector with entries $V_n^{(\ell)} = \frac{(d-1)!(dn-2T_{\ell})}{\ell!(\ln(n))^{d-1-\ell}}$ $\frac{a-1 \cdot \ldots \cdot (an-21 \ell)}{\ell!(\ln(n))^{d-1-\ell}}$ converges in distribution to the product of *d* − 1 independent exponential variables of mean 1. \Box

The rest of this section is for proving the first statement of Theorem [8](#page-7-2) (for some fixed *k* using induction), and Section [5](#page-13-0) will be for proving the second statement (again, for some fixed *k* using induction, assuming the first statement holds for the same *k*). Hence, for the rest of the paper we will fix some $k \in \{0\}$ ∪ $[d-2]$.

First, we note that [\(16\)](#page-7-3) holds w.h.p. for all *j* < *k* by a simple argument: by induction on the second statement of Theorem [8,](#page-7-2) w.h.p. if $i \in I_k$ then $S_i^{(j)} = 0$. By [\(11\)](#page-4-5) and by definition of $E_j(i)$, if *i* ∈ *I_k* then $ns_i(i/n)$ ≪ $E_i(i)$, completing the argument.

Next, we prove that [\(16\)](#page-7-3) holds for $j = d - 1$ if it holds for all other values of *j*: by combining [\(1\)](#page-2-1) and [\(7\)](#page-3-2), we have

$$
\left| S_i^{(d-1)} - ns_{d-1} \left(\frac{i}{n} \right) \right| = \left| \sum_{j=0}^{d-2} \left(S_i^{(j)} - ns_j \left(\frac{i}{n} \right) \right) \right|
$$

$$
\leq \sum_{j=0}^{d-2} 4E_j(i) \qquad \text{by (16) for } j \leq d-2
$$

$$
< 4E_{d-1}(i) \qquad \text{by (15).}
$$

Hence, for the rest of this section, we need to show the first statement of Theorem [8](#page-7-2) for $j \in$ [*k*, *d* − 2]. From now on we always assume *j* to be in this range. We will *also* assume that, for all $\lambda < k$, $S_i^{(\lambda)} = 0$ if $i \in I_k$ (which holds w.h.p. from above).

In this section we will make use of so-called *critical intervals*, ranges of possible values for $S_i^{(j)}$ *i* in which we apply a martingale argument. The lower critical interval will be

$$
[ns_j(i/n) - 4E_j(i), ns_j(i/n) - 3E_j(i)],
$$

and the upper critical interval will be

$$
[ns_j(i/n) + 3E_j(i), ns_j(i/n) + 4E_j(i)].
$$

Our goal is to show that w.h.p. $S_i^{(j)}$ does not cross either critical interval; however, we first need to show that $S_i^{(j)}$ sits between the critical intervals at the beginning of the phase (this is the reason why *Ej*(*i*) has the 2*^k* factor; it makes a sudden jump in size between phases to accommodate a new martingale process), which is the statement of our first Lemma of this section:

Lemma 9. *W.h.p., for all j* ∈ [*k, d* − 2] *(putting i_{after}*(−1) = *i_{trans} for convenience of notation)*:

$$
\left|S_{i_{after}(k-1)+1}^{(j)}-ns_j\left(\frac{i_{after}(k-1)+1}{n}\right)\right| < 3E_j(i_{after}(k-1)+1).
$$

Proof. First, recall that $S_{i+1}^{(j)} - S_i^{(j)} \in \{-2, -1, 0\}$ and $|ns_j((i+1)/n) - ns_j(i/n)| \le 2$ for any *i* and *j* (see paragraph above (14)). Second, consider the change in the bound itself between $i_{after}(k-1)$ and $i_{after}(k-1) + 1$: by definitions of i_{trans} , E_{first} , E_j , and by [\(15\)](#page-7-4), we have $1 \ll E_{first}(i_{trans}) =$ $\Theta(n^{0.64})$, $E_j(i_{trans} + 1) = \omega(n^{0.69})$, and $1 \ll E_j(i_{after}(k-1)) \approx \frac{1}{2}(E_j(i_{after}(k-1) + 1)$ for $k > 0$. Hence, by induction on the first statement of Theorem [8](#page-7-2) and by Theorem [6,](#page-5-2) the statement of the Lemma follows.

Next, like in Section [3,](#page-5-0) we define two new random variables for each *j* and $i \in I_k$:

$$
S_i^{(j)+} := S_i^{(j)} - ns_j(i/n) - 4E_j(i)
$$

$$
S_i^{(j)-} := S_i^{(j)} - ns_j(i/n) + 4E_j(i).
$$

We also re-introduce the stopping time *T*, now defined as the first step $i \in I_k$ for which [\(16\)](#page-7-3) is *not* satisfied for some *j*; if [\(16\)](#page-7-3) always holds, then let $T = \infty$. Let variable name *W* be introduced to equip this stopping time to variable *S*, i.e.

$$
W_i^{(j)+} := \begin{cases} S_i^{(j)+}, i < T \\ S_T^{(j)+}, i \ge T \end{cases} \qquad W_i^{(j)-} := \begin{cases} S_i^{(j)-}, i < T \\ S_T^{(j)-}, i \ge T. \end{cases}
$$

Note that $W_i^{(j)+}$ corresponds to the upper critical interval, and $W_i^{(j)-}$ to the lower one. $Further more, the inequality of Theorem 8 holds if and only if $W_{i_{after}(k)}^{(j)+} \leq 0$ and $W_{i_{after}(k)}^{(j)-} \geq 0$ for$ $Further more, the inequality of Theorem 8 holds if and only if $W_{i_{after}(k)}^{(j)+} \leq 0$ and $W_{i_{after}(k)}^{(j)-} \geq 0$ for$ $Further more, the inequality of Theorem 8 holds if and only if $W_{i_{after}(k)}^{(j)+} \leq 0$ and $W_{i_{after}(k)}^{(j)-} \geq 0$ for$ each *j* (here we must make use of our assumption that $S_i^{(\lambda)} = 0$ for all $\lambda < k$). The next Lemma states that, within their respective critical intervals, they are a supermartingale and submartingale respectively:

Lemma 10. For all
$$
i \in I_k
$$
 and for all $j \in [k, d-2]$, $\mathbb{E}\left[W_{i+1}^{(j)-} - W_i^{(j)-} | G_i\right] \ge 0$ whenever $W_i^{(j)-} \le E_j(i)$, and $\mathbb{E}\left[W_{i+1}^{(j)+} - W_i^{(j)+} | G_i\right] \le 0$ whenever $W_i^{(j)+} \ge -E_j(i)$.

Proof. Here we just prove the first part of the Lemma; the second part follows from nearly identical calculations. By the same logic as in the proof of Lemma [7](#page-5-3) we work with *S*(*j*)[−] instead of *W*(*j*)[−] and assume that [\(16\)](#page-7-3) holds for all *j*. We also have the same expected change as in Lemma [7,](#page-5-3) except with $E_{first}(i)$ replaced with $4E_j(i)$:

$$
\mathbb{E}\left[S_{i+1}^{(j)-} - S_i^{(j)-} | G_i\right] = \frac{-2Y_i^{(j)}}{S_i^{(d-1)}} + O\left(\frac{1}{dn-2i}\right) + \frac{2y_j(i/n)}{s_{d-1}(i/n)} - \frac{d^2}{d\mu^2} \left(\frac{ns_j(\mu/n)}{2}\right)\Big|_{\mu=\psi} + 4(E_j(i+1) - E_j(i)).
$$

We split the above expression $\left(\text{excluding } O\left(\frac{1}{dn-2i}\right)\right)$ into three summands, assuming $S_i^{(j)-} \leq$ $E_j(i) \iff S_i^{(j)} - ns_j(i/n) \le -3E_j(i)$ (for convenience, for the case *j* = 0, we put $S_i^{(j-1)}$, *s_{j−1}*, and *Ej*[−]¹ all equal to 0 :

1.

$$
\frac{-2Y_i^{(j)}}{S_i^{(d-1)}} + \frac{2y_j(i/n)}{s_{d-1}(i/n)} = \frac{-2S_i^{(j)} + 2ns_j(i/n) + 2S_i^{(j-1)} - 2ns_{j-1}(i/n)}{ns_{d-1}(i/n)}
$$

+ $2Y_i^{(j)} \left(\frac{1}{ns_{d-1}(i/n)} - \frac{1}{S_i^{(d-1)}}\right)$

$$
\geq \frac{6E_j(i) - 8E_{j-1}(i)}{ns_{d-1}(i/n)} - \frac{8S_i^{(j)}E_{d-1}(i)}{S_i^{(d-1)}(ns_{d-1}(i/n))} \qquad \text{by (16)}
$$

$$
\geq \frac{5.9E_j(i)}{ns_{d-1}(i/n)} - \frac{9S_i^{(j)}E_{d-1}(i)}{(ns_{d-1}(i/n))^2} \qquad \text{by (15), (16), (11), and } i \leq i_{after}(k)
$$

$$
\geq \frac{5.9E_j(i)}{ns_{d-1}(i/n)} - \frac{9(ns_j(i/n))E_{d-1}(i)}{(ns_{d-1}(i/n))^2} - \frac{36E_j(i)E_{d-1}(i)}{(ns_{d-1}(i/n))^2} \qquad \text{by (16)}
$$

$$
= \left(\frac{E_j(i)}{ns_{d-1}(i/n)}\right) \left(5.9 - 9\left(\frac{s_j(i/n)}{s_{d-1}(i/n)}\right)^{0.3} - \frac{36 * 2^k \ln(n)^{0.05}}{(ns_{d-1}(i/n))^{0.3}}\right)
$$

$$
\geq \frac{5.8E_j(i)}{ns_{d-1}(i/n)} \qquad \text{by } i \leq i_{after}(k) \text{ and (11)}.
$$

2. Just as in the proof of Lemma [7:](#page-5-3)

$$
-\frac{d^2}{d\mu^2}\left(\frac{ns_j(\mu/n)}{2}\right)\Big|_{\mu=\psi}=O\left(\frac{1}{dn-2i}\right).
$$

3.

$$
4(E_j(i+1) - E_j(i)) = 4 \frac{dE_j(\mu)}{d\mu}\Big|_{\mu = \phi} \quad \text{for some } \phi \in [i, i+1]
$$

= (4)(2^k) ln (n)^{0.05} $\left(\frac{0.7}{(ns_j(\phi/n))^{0.3}}\right) \left(\frac{-2y_j(\phi/n)}{s_{d-1}(\phi/n)}\right) \quad \text{by (5)}$
= (4 + o(1))(2^k) ln (n)^{0.05} $\left(\frac{0.7}{(ns_j(\phi/n))^{0.3}}\right) \left(\frac{-2s_j(\phi/n)}{s_{d-1}(\phi/n)}\right) \quad \text{by (10)}$
= $\frac{-(5.6 + o(1))E_j(\phi)}{ns_{d-1}(\phi/n)} = \frac{-(5.6 + o(1))E_j(i)}{ns_{d-1}(i/n)}.$ (17)

Now we put the above bounds together (using [\(11\)](#page-4-5), [\(15\)](#page-7-4), and $i \leq i_{after}(k) \leq i_{after}(j)$:

$$
\mathbb{E}\left[S_{i+1}^{(j)-} - S_i^{(j)-} | G_i\right] \ge \frac{0.01E_j(i)}{ns_{d-1}(i/n)} + O\left(\frac{1}{dn-2i}\right) \ge 0.
$$

 \Box

Figure 1. Visual representation of event $\mathcal{E}_{\ell}^{(\bar{y})+}.$

We introduce the next Lemma to get sufficiently small bounds on the one-step changes in each time step (this is known as the "bounded hypothesis" from [\[14\]](#page-18-5)):

Lemma 11. *For all i* ∈ I_k *and all j* ∈ $[k, d-2]$ *,*

$$
-3 < W_{i+1}^{(j)\xi} - W_i^{(j)\xi} < \ln\left(n\right)^{-d+1.06+j}
$$

*where "*ξ *" can be either "*+*" or "*−*".*

Proof. Like in the proofs of Lemma [7](#page-5-3) and [10,](#page-9-0) we assume that $W^{\xi} = S^{(j)\xi}$ (ξ is + or −), else $W_{i+1}^{\xi} - W_i^{\xi} = 0$. Again, we have $-2 \le S_{i+1}^{(j)} - S_i^{(j)} \le 0$. Secondly, we have

$$
|-ns_j((i+1)/n) + ns_j(i/n) - CE_j(i+1) + CE_j(i)|
$$

\n
$$
\leq |-ns_j((i+1)/n) + ns_j(i/n)| + |-CE_j(i+1) + CE_j(i)|
$$

\n
$$
= O\left(\frac{y_j(i/n)}{s_{d-1}(i/n)} + \frac{E_j(i)}{ns_{d-1}(i/n)}\right) \qquad \text{by (5), (11), and (17)}
$$

\n
$$
= o\left(\ln(n)^{-d+1.06+j}\right) \qquad \text{by (11), (15), and } i \leq i_{after}(k) \leq i_{after}(j).
$$

Combining the inequalities completes the proof. \Box

To put this all together to prove the first part of Theorem [8,](#page-7-2) we introduce a series of events: first, let $\mathcal{E}^{(j)+}$ denote the event that $W^{(j)+}_{i_{after}(k)} > 0$ and $\mathcal{E}^{(j)-}$ denote the event that $W^{(j)-}_{i_{after}(k)} < 0$. Let $\mathcal{E} =$ $\left(\bigcup_{j\geq k} \mathcal{E}^{(j)+}\right)\cup\left(\bigcup_{j\geq k} \mathcal{E}^{(j)-}\right)$; we seek to bound $\mathbb{P}[\mathcal{E}]$, since $\mathcal E$ is the event that [\(16\)](#page-7-3) *doesn't* hold for some $i \in I_k$. Next, for all $\ell \in I_k$, let $\mathcal{H}^{(j)+}_\ell$ be the event that $W^{(j)+}_{\ell-1} < -E_j(\ell-1)$ and $W^{(j)+}_\ell \geq -E_j(\ell)$, and let

$$
\mathcal{E}_{\ell}^{(j)+} := \mathcal{H}_{\ell}^{(j)+} \cap \left\{ W_{i}^{(j)+} \geq -E_{j}(i) \text{ for all } i \geq \ell \right\} \cap \left\{ W_{i_{after(k)}}^{(j)+} > 0 \right\}.
$$

(see Fig. [1](#page-11-0) for a visual representation of event $\mathcal{E}_{\ell}^{(j)+}\Big)$

Similarly, for all $\ell \in I_k$, let $\mathcal{H}^{(j)-}_\ell$ be the event that $W^{(j)-}_{\ell-1} > E_j(\ell-1)$ and $W^{(j)-}_\ell \leq E_j(\ell)$, and let

$$
\mathcal{E}_{\ell}^{(j)-} := \mathcal{H}_{\ell}^{(j)-} \cap \left\{ W_{i}^{(j)-} \leq E_j(i) \text{ for all } i \geq \ell \right\} \cap \left\{ W_{i_{after}(k)}^{(j)-} < 0 \right\}.
$$

Finally, note that, by Lemma [9,](#page-9-1) with high probability we must have

$$
W_{i_{after}(k-1)+1}^{(j)+} < -E_j(i_{after}(k-1)+1) \text{ and } W_{i_{after}(k-1)+1}^{(j)-} > E_j(i_{after}(k-1)+1).
$$

Furthermore, assuming these two inequalities hold (and, once again, assuming that $S_i^{\lambda} = 0$ if λ *k*), then if $W_{i_{after}(k)}^{(j)+} > 0$ for some *j*, one of the events $\mathcal{E}_{\ell}^{(j)+}$ must happen; likewise, if $W_{i_{after}(k)}^{(j)-} < 0$ for some *j*, one of the events $\mathcal{E}_{\ell}^{(j)-}$ must happen; hence, $\mathcal{E}^{(j)+} = \bigcup \mathcal{E}_{\ell}^{(j)+}$ and $\mathcal{E}^{(j)-} = \bigcup \mathcal{E}_{\ell}^{(j)-}$. ℓ

We are now ready to prove the first statement of Theorem [8](#page-7-2) in full.

Proof of the first part of Theorem [8](#page-7-2) *with fixed k*. First, we fix an arbitrary *j* (in $[k, d-2]$). We prove that $\mathbb{P}[\mathcal{E}^{(j)-}] = \exp\big\{-\Omega\big(\ln(n)^{0.036}\big)\big\}$; the proof for bounding $\mathbb{P}[\mathcal{E}^{(j)+}]$ is nearly identical. We will use Corollary [5](#page-4-2) to bound $\mathbb{P}[\mathcal{E}_{\ell}^{(j)-}]$ for each fixed ℓ . Given a fixed ℓ , we define a modified stopping time

$$
T_{mod} := \min_{i \in [\ell, i_{after}(k)]} \left\{ W_i^{(j)-} > E_j(i) \text{ or } i = T \right\}
$$

(letting $T_{mod} = \infty$ if the condition doesn't hold for any *i* in the range). Let variable W_i^{ℓ} be the variable $W_i^{(j)-}$ defined just on $i \in [\ell, i_{after}(k)]$ equipped with this stopping time (we drop the " (j) −" here for convenience); i.e.

$$
W_i^{\ell} := \begin{cases} W_i^{(j)-}, i < T_{mod} \\ W_{T_{mod}}^{(j)-}, i \ge T_{mod}. \end{cases}
$$

Note that $(W_i^{\ell})_i$ (over $i \in [\ell, i_{after}(k)])$ is a submartingale by Lemma [10,](#page-9-0) since our new stopping time negates the need for the condition $W_i^{(j)-}$ ≤ $E_j(i)$; also, $(W_i^{\ell})_i$ satisfies Lemma [11.](#page-11-1) Since we want an upper bound for $\mathbb{P}[\mathcal{E}_{\ell}^{(j)-}]$, we can condition on event $\mathcal{H}_{\ell}^{(j)-}$, as $\mathcal{H}_{\ell}^{(j)-} \supseteq \mathcal{E}_{\ell}^{(j)-}$. Now let

$$
A_i = -W_{\ell+i}^{\ell} + W_{\ell}^{\ell},
$$

\n
$$
\eta = \ln(n)^{-d+1.06+j},
$$

\n
$$
N = 3,
$$

\n
$$
m = i_{after}(k) - \ell,
$$

\n
$$
a = 0.9E_j(\ell).
$$

Note that the conditions of Corollary [5](#page-4-2) are satisfied: $0 = A_0$ and $\eta < N/10$ are obvious, Lemma [11](#page-11-1) gives us $-\eta \leq A_{i+1} - A_i \leq N$, and $(A_i)_i$ is a supermartingale since $(W_i^{\ell})_i$ is a sub-martingale. We therefore implement Corollary [5,](#page-4-2) using $m \leq \frac{dn}{2} - \ell \leq dns_{d-1}(\ell/n)$ (by [\(9\)](#page-4-4)), [\(11\)](#page-4-5), and [\(15\)](#page-7-4):

$$
\mathbb{P}[A_m \ge a] \le e^{-\frac{a^2}{3\eta N m}} + e^{-\frac{a}{6N}} = e^{-\Omega(\ln(n)^{0.04}(ns_j(\ell/n))^{0.4})} + e^{-\Omega(\ln(n)^{0.05}(ns_j(\ell/n))^{0.7})}.
$$
 (18)

To bound $\mathbb{P}[\mathcal{E}_{\ell}^{(j)}]$, we show that $\mathcal{E}_{\ell}^{(j)}$ = { $A_m \ge a$ } and apply [\(18\)](#page-12-0) while conditioning on $\mathcal{H}_{\ell}^{(j)}$. Given $\mathcal{H}_{\ell}^{(j)}$ happens, we have $W_{\ell}^{\ell} = W_{\ell}^{(j)} > 0.9E_j(\ell) = a$ by [\(15\)](#page-7-4), Lemma [11,](#page-11-1) and $i \leq i_{after}(j)$. Therefore $\mathcal{E}_{\ell}^{(j)-} = \mathcal{H}_{\ell}^{(j)-} \cap \left\{ W_{i_{after}(k)}^{\ell} < 0 \right\} \subseteq \{ A_m \geq a \},$ hence

$$
\mathbb{P}\left[\mathcal{E}_{\ell}^{(j)-}\right] = e^{-\Omega\left(\ln(n)^{0.04}(ns_j(\ell/n))^{0.4}\right)} + e^{-\Omega\left(\ln(n)^{0.05}(ns_j(\ell/n))^{0.7}\right)}.
$$

We now take a union bound to bound $\mathbb{P}[\mathcal{E}^{(j)-}]$ (using [\(11\)](#page-4-5) where appropriate):

$$
\mathbb{P}[\mathcal{E}^{(j)}] \leq \sum_{\ell=i_{after}(k-1)+1}^{i_{after}(k)} \mathbb{P}[\mathcal{E}_{\ell}^{(j)} -]
$$
\n
$$
= \sum_{\ell=i_{trans}}^{i_{after}(k)} (\exp \{-\Omega (\ln (n)^{0.04} (n s_j(\ell/n))^{0.4})\} + \exp \{-\Omega (\ln (n)^{0.05} (n s_j(\ell/n))^{0.7})\})
$$
\n
$$
= \sum_{\ell=i_{trans}}^{i_{after}(j)} (\exp \{-\Omega (\frac{(dn - 2\ell)^{0.4}}{\ln (n)^{0.4d - 0.44 - 0.4j}})\} + \exp \{-\Omega (\frac{(dn - 2\ell)^{0.7}}{\ln (n)^{0.7d - 0.75 - 0.7j}})\})
$$
\n
$$
= \sum_{p=\lfloor \ln (n)^{d-1.01-j} \rfloor}^{[n^{1-1/(100d)]}} (\exp \{-\Omega (\frac{p^{0.4}}{\ln (n)^{0.4d - 0.44 - 0.4j}}))\} + \exp \{-\Omega (\frac{p^{0.7}}{\ln (n)^{0.7d - 0.75 - 0.7j}})\})
$$
\n
$$
= \ln (n)^{d-1.01-j} \sum_{q=1}^{\infty} (\exp \{-\Omega (q^{0.4} \ln (n)^{0.036})\} + \exp \{-\Omega (q^{0.7} \ln (n)^{0.043})\})
$$
\n
$$
= \exp \{-\Omega (\ln (n)^{0.036})\}.
$$

We give a note for the aspects of the proof of bounding $\mathbb{P}[\mathcal{E}^{(j)+}]$ that are different from the above: use the variable $W_i^{(j)+}$ instead of $W_i^{(j)-}$, events $\mathcal{E}_\ell^{(j)+}$ instead of $\mathcal{E}_\ell^{(j)-}$, and $\mathcal{H}_\ell^{(j)+}$ instead of *H*(*j*)[−] - . Define *Tmod* instead as

$$
T_{mod} := \min_{i \in [\ell, i_{after}(k)]} \left\{ W_i^{(j)+} < -E_j(i) \text{ or } i = T \right\}.
$$

Finally, use Corollary [4](#page-4-0) instead of Corollary [5](#page-4-2) (which will be slightly easier to implement). \Box

5. Final phase

We continue our proof by induction of Theorem [8](#page-7-2) with our fixed index *k*; now we prove the second part. We assume the first part of Theorem [8](#page-7-2) to hold, as well as the second part of the Theorem for lesser *k*; for example, we have $S_{i_{after}(k-1)}^{(k-1)} = 0$ w.h.p. In this section we focus on the *d*-process for a narrow domain of *i*. Let

$$
i_{before}(k) := \left\lfloor \frac{dn}{2} - \ln{(n)}^{d-0.8-k} \right\rfloor.
$$

We will consider the *d*-process starting at step $i_{before}(k)$ assuming that [\(16\)](#page-7-3) holds at $i = i_{before}(k)$; we do not need the first part of Theorem [8](#page-7-2) in this section otherwise. We do not use martingale arguments here, but rather we show that the distribution of the sequence of time steps at which a vertex of degree *k* is chosen from the *d*-process is similar to a uniform distribution over all pos-sible such sequences. Theorem [8,](#page-7-2) [\(10\)](#page-4-6), and [\(15\)](#page-7-4) tell us that w.h.p. we will have $\sim \frac{2(d-1)!}{k!} \ln(n)^{0.2}$ vertices of degree at most *k* (or degree equal to *k*; they are the same here) left when there are [ln (*n*)^{*d*−0.8−*k*}] steps left; hence, the average distance between steps at which we remove vertices of degree *k* is $\frac{k!}{2(d-1)!} \ln(n)^{d-1-k}$. When there are this many steps left times *r*, we expect *r* such

vertices to remain, and for the probability that there are no vertices of degree *k* to be *e*−*^r* . Most of this section will build towards proving the following Theorem:

Theorem 12. Let $L(n)$ be an integer-valued function so that $L(n) = \Theta(\ln(n)^{0.2})$ and let $J(n) =$ $\lfloor \frac{dn}{2} \rfloor$ − *i_{before}*(*k*) ∼ ln (*n*)^{*d*−0.8−*k*. Let H be any graph with *i*_{before}(*k*) edges which satisfies [\(16\)](#page-7-3) at *i* =} $i_{before}(k)$, has no vertices of degree at most $k-1$, and has $L(n)$ vertices of degree k. Also, let $r \in \mathbb{R}^+$ *be arbitrary. Then*

$$
\mathbb{P}\left[S_{\lfloor \frac{dn}{2} - \frac{rJ(n)}{L(n)}\rfloor}^{(k)} = 0 \middle| G_{i_{before}(k)} = H \right] \to e^{-r}.
$$

First, we note that, given that (16) holds for $i = i_{before}(k)$ and by (1) , that $w.h.p.$ *dn* − 2*i*_{before}(*k*) − $S^{(d-1)}_{i_{before}(k)} = O\left(\frac{dn-2i_{before}}{\ln(n)}\right)$ (consider $S^{(d-2)}_{i_{before}(k)}$ $\binom{(d-2)}{i_{before(k)}}$; hence, for all $i \in [i_{before}(k), i_{after}(k)]$:

$$
S_i^{(d-1)} = dn - 2i + O\left(\frac{dn - 2i_{before}}{\ln(n)}\right) = (dn - 2i)\left(1 + O\left(\frac{1}{\ln(n)^{0.79}}\right)\right). \tag{19}
$$

Let $t_{start} = i_{before}(k)$ and $t_{end} = \lfloor dn/2 - rJ(n)/L(n) \rfloor$. Consider the *d*-process between t_{start} and t_{end} , given that $\tilde{G}_{t_{start}} = H$. At each step two vertices are chosen; now assume that the pair at each step is ordered uniformly at random, so that a sequence of 2(*tend* − *tstart*) vertices is generated. We also generate a *binary* sequence simultaneously, each digit corresponding to a vertex: after a pair of vertices is picked for the vertex sequence, for each of the two vertices (in the order that they are randomly shuffled) append a "1" to the binary sequence if the corresponding vertex had degree *k* just before it was picked, and append a "0" otherwise. Let $P:\{0, 1\}^{2(t_{end}-t_{start})} \to [0, 1]$ be the corresponding probability function that arises from this process (note that, if γ is a string with more than *L*(*n*) "1"'s, then $P(\gamma) = 0$). Note that P depends on the graph *H*. We compare this to a second probability function Q :{0, 1}^{2(t_{end} - t_{start}) \rightarrow [0, 1], which is defined by picking a binary} string with $L(n)$ 1's and $2J(n) - L(n)$ 0's uniformly at random, then taking the first $2(t_{end} - t_{start})$ digits.

For any binary sequence γ with ℓ digits, and $I \subset [\ell]$, let γ_I be the subsequence with indices from *I*; for example, $\gamma_{[a]}$ would be the first *a* digits of γ , and $\gamma_{[a]}$ would just be the *a*-th digit (for notation's sake, let " $\gamma_{[0]}$ " be the empty string). Also let $\|\gamma\|$ denote the number of 1's in γ . We now present the following Lemma:

Lemma 13. *Let* α *be an arbitrary* $2(t_{end} - t_{start})$ *length binary sequence with at most* $L(n)$ *1's, and let* γ *be the random binary sequence according to either P or Q. Let i* ∈ [2(*tend* − *tstart*)]*. Then (letting* $a_{\{0\}} = 1$ *for sake of notation*):

$$
\frac{\mathbb{P}_{\mathcal{D}}[\gamma_{[i]} = \alpha_{[i]} \mid \gamma_{[i-1]} = \alpha_{[i-1]}]}{\mathbb{P}_{\mathcal{Q}}[\gamma_{[i]} = \alpha_{[i]} \mid \gamma_{[i-1]} = \alpha_{[i-1]}]} \begin{cases}\n= 1 + O\left(\frac{1}{J(n) \ln(n)^{0.39}}\right) & if \alpha_{\{i\}} = 0 \text{ and } \alpha_{\{i-1\}} = 0 \\
= 1 + O\left(\frac{\ln(n)^{0.4}}{J(n)}\right) & if \alpha_{\{i\}} = 0 \text{ and } \alpha_{\{i-1\}} = 1 \\
= 1 + O\left(\frac{1}{\ln(n)^{0.79}}\right) & if \alpha_{\{i\}} = 1 \text{ and } \alpha_{\{i-1\}} = 0 \\
\leq 1 + O\left(\frac{1}{\ln(n)^{0.79}}\right) & if \alpha_{\{i\}} = 1 \text{ and } \alpha_{\{i-1\}} = 1.\n\end{cases}
$$

Proof. First, we consider the cases where $\alpha_{\{i\}} = 1$. We have

$$
\mathbb{P}_{\mathcal{Q}}[\gamma_{\{i\}}=1 \mid \gamma_{[i-1]}=\alpha_{[i-1]}]=\frac{L(n)-\|\alpha_{[i-1]}\|}{2J(n)-(i-1)}.
$$
\n(20)

For the probability space *P*, we need to consider three subcases: we need to consider whether *i* is even or odd, and if it is even, whether α_{i-1} is 0 or 1, since each step of the *d*-process outputs two digits of the binary string. Let's say that τ corresponds to the last step in the *d*-process before the *i*-th binary digit is generated (recall that pairs of digits are generated together). Then if *i* is odd:

$$
\mathbb{P}_{\mathcal{P}}[\gamma_{\{i\}} = 1 \mid \gamma_{[i-1]} = \alpha_{[i-1]}] = -\frac{1}{2} \mathbb{E} \left[S_{\tau+1}^{(k)} - S_{\tau}^{(k)} | G_{\tau} \right]
$$

\n
$$
= \frac{S_{\tau}^{(k)}}{S_{\tau}^{(d-1)}} \left(1 + O\left(\frac{1}{dn - 2\tau} \right) \right) \qquad \text{by (3)}
$$

\n
$$
= \frac{S_{\tau}^{(k)}}{2 \lfloor dn/2 - \tau \rfloor} \left(1 + O\left(\frac{1}{\ln(n)^{0.79}} \right) \right) \qquad \text{by (19)}
$$

\n
$$
= \frac{L(n) - ||\alpha_{[i-1]}||}{2J(n) - (i-1)} \left(1 + O\left(\frac{1}{\ln(n)^{0.79}} \right) \right). \qquad (21)
$$

If *i* is even and $\alpha_{\{i-1\}} = 1$, then $S_{\tau}^{(k)} = L(n) - ||\alpha_{\{i-1\}}|| + 1$. At step τ there are $S_{\tau}^{(k)}\left(S_{\tau}^{(d-1)}+O(1)\right)$ ordered pairs of vertices whose first vertex has degree *k*, and *at most* 2 $\binom{S_{\tau}^{(k)}}{2}$ ordered pairs of vertices both with degree *k*; hence:

$$
\mathbb{P}_{\mathcal{P}}[\gamma_{\{i\}} = 1 \mid \gamma_{[i-1]} = \alpha_{[i-1]}] \le \frac{S_{\tau}^{(k)} - 1}{S_{\tau}^{(k-1)} + O(1)} \n= \frac{S_{\tau}^{(k)} - 1}{2\lfloor dn/2 - \tau \rfloor} \left(1 + O\left(\frac{1}{\ln(n)^{0.79}}\right)\right) \qquad \text{by (19)} \n= \frac{L(n) - ||\alpha_{[i-1]}||}{2J(n) - (i-1)} \left(1 + O\left(\frac{1}{\ln(n)^{0.79}}\right)\right).
$$
\n(22)

Hence the final inequality of the Lemma holds by (21) and (22) .

Next, consider the case where *i* is even and $\alpha_{\{i-1\}} = 0$; here, $S_{\tau}^{(k)} = L(n) - ||\alpha_{[i-1]}||$ once again. At step τ there are $\left(S_{\tau}^{(d-1)}-S_{\tau}^{(k)}\right)\left(S_{\tau}^{(d-1)}+O(1)\right)$ ordered pairs of vertices whose first vertex has degree greater than k , and $S_{\tau}^{(k)}\left(S_{\tau}^{(d-1)}-S_{\tau}^{(k)}+O(1)\right)$ ordered pairs of vertices for which the first vertex has degree greater *k* and the second vertex has degree *k* (one can "pick the second vertex first" to see this). Hence:

$$
\mathbb{P}_{\mathcal{P}}[\gamma_{\{i\}} = 1 \mid \gamma_{[i-1]} = \alpha_{[i-1]}\n = \frac{S_{\tau}^{(k)}}{S_{\tau}^{(d-1)}} \left(1 + O\left(\frac{1}{S_{\tau}^{(d-1)}}\right)\right) \quad \text{since } S_{\tau}^{(d-1)} \gg S_{\tau}^{(k)} \text{ there}
$$
\n
$$
= \frac{S_{\tau}^{(k)}}{2\lfloor d n/2 - \tau \rfloor} \left(1 + O\left(\frac{1}{\ln(n)^{0.79}}\right)\right) \quad \text{by (19)}
$$
\n
$$
= \frac{L(n) - ||\alpha_{[i-1]}||}{2J(n) - (i-1)} \left(1 + O\left(\frac{1}{\ln(n)^{0.79}}\right)\right), \quad (23)
$$

hence the third equality of the Lemma holds by (21) and (23) .

Now consider $\alpha_{\{i\}} = 0$. By modifying [\(20\)](#page-14-0) to accommodate $\gamma_{\{i\}} = 0$, we have

$$
\mathbb{P}_{\mathcal{Q}}[\gamma_{\{i\}}=0 \mid \gamma_{[i-1]}=\alpha_{[i-1]}\rbrack=1-\frac{L(n)-\|\alpha_{[i-1]}\|}{2J(n)-(i-1)}.
$$
\n(24)

Similarly, by modifying [\(21\)](#page-15-0) and [\(23\)](#page-15-2), if $a_{\{i-1\}} = 0$ then

$$
\mathbb{P}_{\mathcal{P}}[\gamma_{\{i\}}=0 \mid \gamma_{[i-1]}=\alpha_{[i-1]}]=1-\frac{L(n)-\|\alpha_{[i-1]}\|}{2J(n)-(i-1)}\left(1+O\left(\frac{1}{\ln(n)^{0.79}}\right)\right). \tag{25}
$$

By modifying [\(21\)](#page-15-0) and [\(22\)](#page-15-1), if $a_{\{i-1\}} = 1$, then

$$
\mathbb{P}_{\mathcal{P}}[\gamma_{\{i\}} = 0 \mid \gamma_{[i-1]} = \alpha_{[i-1]}\] \ge 1 - \frac{L(n) - ||\alpha_{[i-1]}||}{2J(n) - (i-1)} \left(1 + O\left(\frac{1}{\ln(n)^{0.79}}\right)\right)
$$

$$
= 1 + O\left(\frac{L(n) - ||\alpha_{[i-1]}||}{2J(n) - (i-1)}\right). \tag{26}
$$

Since $L(n) = \Theta(\ln(n)^{0.2})$ and

$$
2J(n) - (i - 1) = 2\lfloor dn/2 - \tau \rfloor = \Omega(dn/2 - t_{end}) = \Omega(J(n)/\ln(n)^{0.2}),
$$

then $\frac{L(n)-\|\alpha_{i-1}\|}{2J(n)-(i-1)} = O\left(\frac{\ln(n)^{0.4}}{J(n)}\right)$ $\frac{(n)^{0.4}}{J(n)}$. Therefore the ratio of [\(25\)](#page-15-3) and [\(24\)](#page-15-4) is $1 + O\left(\frac{1}{J(n)\ln(n)^{0.39}}\right)$, veri-fying the first inequality of the Lemma, and the ratio of [\(26\)](#page-16-0) and [\(24\)](#page-15-4) is $1 + O\left(\frac{\ln(n)^{0.4}}{l(n)}\right)$ $\frac{J(n)^{0.4}}{J(n)}$), verifying the second inequality of the Lemma.

Proof of Theorem [12.](#page-14-1) First, let α be an arbitrary string which satisfies the criteria in Lemma [13.](#page-14-2) By using the Lemma [13](#page-14-2) recursively:

$$
\frac{\mathbb{P}_{\mathcal{P}}[\gamma = \alpha]}{\mathbb{P}_{\mathcal{Q}}[\gamma = \alpha]} \le \exp \left\{ O\left(J(n) \frac{1}{J(n) \ln(n)^{0.39}} + L(n) \left(\frac{1}{\ln(n)^{0.79}} + \frac{\ln(n)^{0.4}}{J(n)} \right) \right) \right\}
$$
\n
$$
= 1 + o(1), \tag{27}
$$

and if α is an arbitrary string *with no two consecutive 1's* which satisfies the criteria in Lemma [13,](#page-14-2) then by similar logic,

$$
\frac{\mathbb{P}_{\mathcal{P}}[\gamma = \alpha]}{\mathbb{P}_{\mathcal{Q}}[\gamma = \alpha]} = 1 + o(1). \tag{28}
$$

Let *C* be the event that γ has two consecutive 1's; we consider $\mathbb{P}[\mathcal{C} | G_{t_{start}} = H]$. We consider probability space *Q* first. Recall that α is a string that has $\sim 2J(n) = \Omega(\ln(n)^{1.2})$ characters and at most $L(n) = \Theta(\ln(n)^{0.2})$ 1's. Because *Q* is a truncation of a uniform distribution, the probability of having two consecutive 1's will be $O\left(\frac{(L(n))^2}{J(n)}\right) = O(\ln(n)^{-0.8})$. Hence, by [\(27\)](#page-16-1) we must have

$$
\mathbb{P}_{\mathcal{P}}[\mathcal{C} \mid G_{t_{start}} = H] = o(1) \text{ and } \mathbb{P}_{\mathcal{Q}}[\mathcal{C} \mid G_{t_{start}} = H] = o(1). \tag{29}
$$

We now combine [\(28\)](#page-16-2) and [\(29\)](#page-16-3) to prove Theorem [12](#page-14-1) (for ease of notation, assume we are given $G_{t_{start}} = H$:

$$
\mathbb{P}\left[S_{\lfloor \frac{d_n}{2} - \frac{r(n)}{L(n)}}^{(k)}\right] = 0\right] = \mathbb{P}_{\mathcal{P}}[\|\gamma\| = L(n) | C] \mathbb{P}_{\mathcal{P}}[C] + \mathbb{P}_{\mathcal{P}}[\|\gamma\| = L(n) | \overline{C}] \mathbb{P}_{\mathcal{P}}[\overline{C}]
$$

\n
$$
= \mathbb{P}_{\mathcal{Q}}[\|\gamma\| = L(n) | C] \mathbb{P}_{\mathcal{P}}[C] + \mathbb{P}_{\mathcal{P}}[\|\gamma\| = L(n) | \overline{C}] \mathbb{P}_{\mathcal{P}}[\overline{C}]
$$

\n
$$
= \frac{\binom{2(t_{end} - t_{start})}{\binom{2f(n)}{L(n)}} + o(1)
$$

\n
$$
= \frac{\binom{2J(n) - 2(\lfloor r/(n)/L(n) \rfloor)}{L(n)}}{\binom{2J(n)}{L(n)}} + o(1)
$$

\n
$$
= \left(1 - \frac{r(1 + o(1))}{L(n)}\right)^{L(n)} + o(1)
$$

\n
$$
= e^{-r} + o(1).
$$

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We can now complete the proof of the second statement of Theorem [8](#page-7-2) at value *k*. Roughly speaking, we will use Theorem [12](#page-14-1) with $L(n) \approx \frac{2(d-1)! \ln(n)^{0.2}}{k!}$, so $\frac{dn}{2} - i(r_k, k) \approx \frac{r_k J(n)}{L(n)}$. First, note that $S_{i_{after}(k)}^{(k)} = 0$ (w.h.p.) comes automatically when the rest of the statement is proved (by putting $r_\ell = 0$ for $\ell < k$ and having $r_k \to 0$). Let \mathcal{G}_ℓ be the event that $\mathcal{S}^{(\ell)}_{\lfloor i(r_\ell,\ell) \rfloor} = 0$ and $\mathcal{G} = \bigcap_{\ell \leq k} \mathcal{G}_\ell$, let \mathcal{F} be the event that [\(16\)](#page-7-3) holds for $i = i_{before}(k)$ and $S_{i_{before}(k)}^{(j)} = 0$ for $j < k$, and let $\mathcal{A} = \mathcal{F} \cap \bigcap_{\ell < k} \mathcal{G}_{\ell}$. Also, let

$$
\mathcal{I} = [ns_k(i_{before}(k)/n) - 4E_k(i_{before}(k)), ns_k(i_{before}(k)/n) + 4E_k(i_{before}(k))].
$$

Note that, by part 1 of Theorem [8,](#page-7-2) by induction on the second part Theorem [8,](#page-7-2) and since $i_{before}(k) > i_{after}(k-1)$, *F* happens with probability $1 - o(1)$. Therefore:

$$
\mathbb{P}[\mathcal{G}] = \mathbb{P}[\mathcal{G}_k \cap \mathcal{A}] + o(1)
$$

= $\sum_{p \in \mathcal{I}} \mathbb{P}\left[\mathcal{G} \middle| \mathcal{A} \cap \left(S_{i_{before}(k)}^{(k)} = p\right)\right] \mathbb{P}\left[\mathcal{A} \cap \left(S_{i_{before}(k)}^{(k)} = p\right)\right] + o(1).$

We can now apply (10) , (15) , and Theorem [12](#page-14-1) to get

$$
\mathbb{P}\left[\mathcal{G} \mid \mathcal{A} \cap \left(S_{i_{before}(k)}^{(k)} = p\right)\right] = e^{-r_{\ell}} + o(1)
$$

for $p \in \mathcal{I}$. We note that all $o(1)$ functions in the sum can be made to be the same by carefully reviewing the proof of Theorem [12.](#page-14-1) Therefore:

$$
\mathbb{P}[\mathcal{G}] = \sum_{p \in \mathcal{I}} (e^{-r_{\ell}} + o(1)) \mathbb{P} \left[\mathcal{A} \cap \left(S_{i_{before}(k)}^{(k)} = p \right) \right] + o(1)
$$

\n
$$
= e^{-r_{\ell}} \sum_{p \in \mathcal{I}} \mathbb{P} \left[\mathcal{A} \cap \left(S_{i_{before}(k)}^{(k)} = p \right) \right] + o(1)
$$

\n
$$
= e^{-r_{\ell}} \mathbb{P} \left[\bigcap_{\ell < k} \mathcal{G}_{\ell} \right] + o(1) \qquad \text{by Theorem 8}
$$

\n
$$
= \exp \left\{ \sum_{\ell=0}^{k} e^{-r_{\ell}} \right\} + o(1) \qquad \text{by induction on Theorem 8,}
$$

proving Theorem [8.](#page-7-2)

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References

- [1] Alon, N. and Spencer, J. (2016) *The Probabilistic Method*. John Wiley & Sons.
- [2] Bohman, T. (2009) The triangle-free process. *Adv. Math.* **221**(5) 1653–1677.
- [3] Bohman, T., Frieze, A. and Lubetzky, E. (2010) A note on the random greedy triangle-packing algorithm. *J. Comb.* **1**(4) 477–488.
- [4] Bohman, T., Frieze, A. and Lubetzky, E. (2015) Random triangle removal. *Adv. Math.* **280** 379–438.
- [5] Bohman, T. and Keevash, P. (2021) Dynamic concentration of the triangle-free process. *Random Struct. Algor.* **58**(2) 221–293.
- [6] Bohman, T. and Picollelli, M. (2012) Sir epidemics on random graphs with a fixed degree sequence. *Random Struct. Algor.* **41**(2) 179–214.
- [7] Janson, S., Ruciński, A. and Luczak, T. (2011) *Random Graphs*. John Wiley & Sons.
- [8] Molloy, M., Surya, E. and Warnke, L. (2022). The degree-restricted random process is far from uniform, arXiv preprint arXiv: [2211.00835.](https://arxiv.org/abs/2211.00835)
- [9] Pontiveros, G. F., Griffiths, S. and Morris, R. (2020) *The Triangle-free Process and the Ramsey Number R(3, k)*. American Mathematical Soc.
- [10] Ruciński, A. and Wormald, N. C. (1997) Random graph processes with maximum degree 2. Ann. Appl. Prob. 7(1) 183–199.
- [11] Rucinski, A. and Wormald, N. C. (1992) Random graph processes with degree restrictions. ´ *Comb. Prob. Comput.* **1**(2) 169–180.
- [12] Ruciński, A. and Wormald, N. (2023). Sharper analysis of the random graph d -process via a balls-in-bins model, arXiv preprint arXiv: [2311.04743.](https://arxiv.org/abs/2311.04743)
- [13] Telcs, A., Wormald, N. and Zhou, S. (2007) Hamiltonicity of random graphs produced by 2-processes. *Random Struct. Algor.* **31**(4) 450–481.
- [14] Wormald, N. C. (1999) The differential equation method for random graph processes and greedy algorithms. *Lect. Approx. Random. Algor.* **73**(155) 0943–05073.

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